Dynamic Portfolio Choice with Linear Rebalancing Rules

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Abstract

We consider a broad class of dynamic portfolio optimization problems that allow for complex models of return predictability, transaction costs, trading constraints, and risk considerations. Determining an optimal policy in this general setting is almost always intractable. We propose a class of linear rebalancing rules, and describe an efficient computational procedure to optimize with this class. We illustrate this method in the context of portfolio execution, and show that it achieves near optimal performance.

1. Introduction

Dynamic portfolio optimization has been a central and essential objective for institutional investors in active asset management. Real world portfolio allocation problems of practical interest have a number of common features:

- **Return predictability.** At the heart of active portfolio management is the fact that a manager will seek to predict future asset returns (Grinold and Kahn, 1999). Such predictions are not limited to simple unconditional estimates of expected future returns. A typical asset manager will make predictions on short- and long-term expected returns using complex models, for example, including return predicting factors such as market capitalization, book-to-market ratio, lagged returns, dividend yields, gross industrial production, and other security specific or macroeconomic variables (see, e.g., Chen et al., 1986; Fama and French, 1996; Goetzmann and Jorion, 1993).

- **Transaction costs.** Trading costs in dynamic portfolio management can arise from sources ranging from the bid-offer spread or execution commissions, to price impact, where the manager’s own trading affects the subsequent evolution of prices. The efficient management of such costs is an important issue broadly, but becomes especially crucial in the setting of optimal
This particular class of portfolio optimization problems seeks to optimally liquidate a given portfolio over a fixed time horizon (Bertsimas and Le, 1998; Almgren and Chriss, 2000).

**Portfolio or trade constraints.** Often times managers cannot make arbitrary investment decisions, but rather face exogenous constraints on their trades or their resulting portfolio. Examples of this include short-sale constraints, leverage constraints, or restrictions requiring market neutrality (or specific industry neutrality).

**Risk aversion.** Portfolio managers seek to control the risk of their portfolios. In practical settings, risk aversion is not accomplished by the specification of an abstract utility function. Rather, managers specify limits or penalties for multiple summary statistics that capture aspects of portfolio risk which are easy to interpret and are known to be important. For example, a manager may both be interested in the risk of the portfolio value changing over various intervals of time, including for example, both short intervals (e.g., daily or weekly risk), as well as risk associated with the terminal value of the portfolio. Such single period risk can be measured a number of ways (e.g., variance, value-at-risk). A manager might further be interested in multiperiod measures of portfolio risk, for example, the maximum drawdown of the portfolio.

Significantly complicating the analysis of portfolio choice is that the underlying problem is multiperiod. Here, in general, the decision made by a manager at a given instant of time might depend on all information realized up to that point. Traditional approaches to multiperiod portfolio choice, dating back at least to the work of Merton (1971), have focused on the analytically determining the optimal dynamic policy. While this work has brought forth important structural insights, it is fundamentally quite restrictive: exact analytical solutions require very specific assumptions investor objectives and market dynamics. These assumptions cannot accommodate flexibility in, for example, the return generating process, trading frictions, and constraints, and are often practically unrealistic. Absent such restrictive assumptions, analytical solutions are not possible. Motivated by this, much of the subsequent academic literature on portfolio choice seeks to develop modeling assumptions that allow for analytical solutions, however the resulting formulations are often not representative of real world problems of practical interest. Further, because of the ‘curse-of-dimensionality’, exact numerical solutions are often intractable as well in cases of practical interest, where the universe of tradeable assets is large.

In search of tractable alternatives, many practitioners eschew multiperiod formulations. Instead, they consider portfolio choice problems in a myopic, single period setting, even when underlying application is clearly multiperiod (e.g., Grinold and Kahn, 1999). Another tractable possibility is to consider portfolio choice problems that are multiperiod, but without the possibility of recourse. Here, a fixed set of deterministic decisions for the entire time horizon are made at the initial time. Both single period and deterministic portfolio choice formulations are quite flexible, and can accommodate many of the features described above. They are typically applied in a quasi-
dynamic fashion through the method of model predictive control. Here, at each time period, the simplified single period or deterministic portfolio choice problem is resolved based on the latest available information. In general, such methods are heuristics; in order to achieve tractability, they neglect the explicit consideration of the possibility of future recourse. Hence, these methods may be significantly suboptimal.

A second tractable alternative is the formulation of portfolio choice problems as a linear quadratic control (e.g., Hora, 2006; Garleanu and Pedersen, 2009). Since at least the 1950’s, linear quadratic control problems have been an important class of tractable multiperiod optimal control problems. In the setting of portfolio choice, if the return dynamics are linear and the transaction costs and risk aversion penalties decomposed into per period quadratic functions, and positions and trading decision are unconstrained, then these methods apply. However, there are many important problem cases that simply do not fall into the linear quadratic framework.

In this paper, our central innovation is to propose a framework for multiperiod portfolio optimization, which admits a broad class of problems including many with features as described earlier. Our formulation maintains tractability by restricting the problem to determining the best policy out of a restricted class of linear rebalancing policies. Such policies allow planning for future recourse, but only of a form that can be parsimoniously parameterized in a specific affine fashion. In particular, the contributions of this paper are as follows:

1. We define a flexible, general setting for portfolio optimization. Our setting allows for very general dynamics of asset prices, which an arbitrary dependence on a history of 'return-predictive factors'. We allow for any convex constraints on trades and positions. Finally, the objective is allowed to be an arbitrary concave function of the sample path of positions. Our framework admits, for example, many complex models for transaction costs or risk aversion.

2. Our portfolio optimization problem is computationally tractable. In our setting, determining the optimal linear rebalancing policy is a convex program. Convexity guarantees that the globally optimal policy can be tractably found in general. This is in contrast to non-convex portfolio choice parametrizations (e.g., Brandt et al., 2009), where only local optimality can be guaranteed.

In our case, numerical solutions can be obtained via, for example, sample average approximation or stochastic approximation methods (see, e.g., Shapiro, 2003; Nemirovski et al., 2009). These methods can be applied in a data-driven fashion, with access only to simulated trajectories and without an explicit model of system dynamics. In a number of instances where the factor and return dynamics are driven by Gaussian uncertainty, we illustrate that our portfolio optimization problem can be reduced to a standard form of convex optimization program, such as a quadratic program or a second order cone program. In such cases, the problem can be solved with off-the-shelf commercial optimization solvers.

3. Our class of linear rebalancing policies subsumes many common heuristic portfolio policies. Both single period and deterministic policies are special cases of linear rebalancing polices,
however linear rebalancing policies are a broader class. Hence, the optimal linear rebalancing policy will outperform policies from these more restricted classes. Further, our method can also be applied in the context of model predictive control. Also, portfolio optimization problems that can be formulated as linear quadratic control also fit in our setting, and their optimal policies are linear rebalancing rules.

4. We demonstrate the practical benefits of our method in an optimal execution example. We consider an optimal execution problem where an investor seeks to liquidate a position. In order to highlight the performance gain using linear decision rules, we use the discrete-time linear quadratic control formulation of Garleanu and Pedersen (2009). However, we further introduce linear inequality constraints that allow the trading decisions to only be sales; such sale-only constraints are common in agency algorithmic trading. We demonstrate that the best linear policy performs better than the best deterministic policy, model predictive control and a projected version of the optimal policy proposed by Garleanu and Pedersen (2009). Further, the performance of the best linear policy is shown to be near optimal, by comparison to upper bounds on optimal policy performance computed for the same problem.

The balance of this paper is organized as follows: In Section 1.1, we review the related literature. In Section 2, we present the abstract form of a dynamic portfolio choice model and provide various specific problems that satisfy the assumptions of the abstract model. We formally describe the class of linear decision rules in Section 3 and discuss solution techniques in order to find the optimal parameters of the linear policy. In Section 4, we provide efficient and exact formulations of dynamic portfolio choice models with Gaussian uncertainty using linear decision models while incorporating linear equality and inequality constraints, proportional and nonlinear transaction costs and a measure of terminal wealth risk. In Section 5, we apply our methodology in an optimal execution problem and evaluate the performance of the best linear policy. Finally, in Section 6, we conclude and discuss some future directions.

1.1. Related Literature

Our paper is related to two different strands of literature: the literature of dynamic portfolio choice with return predictability and transaction costs, and the literature on the use of linear decision rules in the optimal control problems.

First, we consider the literature on dynamic portfolio choice. This vast body of work begins choice starts with the seminal paper of Merton (1971). Following this paper, there has been a significant literature aiming to incorporate the impact of various frictions such as transaction costs on the optimal portfolio choice. For a survey on this literature, see Cvitanic (2001). The work of Constantinides (1986) is an early example that studies the impact of proportional transaction costs on the optimal investment decision and the liquidity premium in the context of CAPM. Davis and Norman (1990), Dumas and Luciano (1991), and Shreve and Soner (1994) provide the exact solution for the optimal investment and consumption decision by formally characterizing the
trade and no-trade regions. One drawback of these papers is that the optimal solution is only computed in the case of a single stock and bond. Liu (2004) extends these result to multiple assets but assumes that asset returns are not correlated.

There is also a significant literature on portfolio optimization that incorporates return predictability (see, e.g., Campbell and Viceira, 2002). Our paper is related to the literature that incorporates both return predictability with transaction costs. Balduzzi and Lynch (1999) and Lynch and Balduzzi (2000) illustrate the impact of return predictability and transaction costs on the utility costs and the optimal rebalancing rule by discretizing the state space of the dynamic program. With a similar state space discretization, Lynch and Tan (2010) model the dynamic portfolio decision with multiple risky assets under return predictability and transaction costs and provide numerical experiments with two risky assets. One significant issue with this line of work is that discretization suffers from the curse-of-dimensionality: the computational effort to determine an optimal policy scales exponentially with the dimension of the state space. When there are more than a few assets or return predicting factors, discretization cannot be applied.

Much of the aforementioned literature seeks to find the best rebalancing policy out of the universe of all possible rebalancing policies. As discussed earlier, this leads to highly restrictive modeling primitives. On the other hand, our work is in the spirit of Brandt et al. (2009), allow for broader modeling flexibility at the expense of considering a restricted class of rebalancing policies. The parameterize the rebalancing rule as a function of security characteristics and estimates the parameters of the rule from empirical data without modeling the distribution of the returns and the return predicting factors. Even though our approach is also a linear parametrization of return predicting factors, there are fundamental differences between our approach and that of Brandt et al. (2009). First, the class of linear polices we consider is much larger than the specific linear functional form in Brandt et al. (2009). In our approach the parameters are time-varying and cross-sectionally different for each security. Second, the extensions provided in Brandt et al. (2009) for imposing positivity constraints and transaction costs are ad-hoc and cannot be generalized to arbitrary convex constraints or transaction cost functions. Finally, the objective function of Brandt et al. (2009) is a non-convex function of the policy parameters. Hence, it is not possible, in general to obtain the globally optimal set of parameters. Our setting, on the other hand, is convex, and hence globally optimal policies can be determined efficiently.

Garleanu and Pedersen (2009) achieve a closed-form solution for a model with linear dynamics in return predictors and quadratic function for transaction costs and quadratic penalty term for risk. However, the analytic solution is highly sensitive to the quadratic cost structure with linear dynamics (see, e.g., Bertsekas, 2000). This special case cannot handle any inequality constraints on portfolio positions, non-quadratic transactions costs, or more complicated risk considerations. On the other hand, our approach can be implemented efficiently in these realistic scenarios and provides more flexibility in the objective function of the investor and the constraints that the investor faces. Boyd et al. (2012) consider an alternative generalization of the linear-quadratic case, using ideas from approximate dynamic programming. Glasserman and Xu (2011) develop a
linear-quadratic formulation for portfolio optimization that offers robustness to modeling errors or mis-specifications.

Second, there is also a literature on the use of linear decision rules in optimal control problems. This approximation technique has attracted considerable interest recently in robust and two-stage adaptive optimization context (see, e.g., Ben-Tal et al., 2004, 2005; Chen et al., 2007, 2008; Bertsimas et al., 2010; Bertsimas and Goyal, 2011). Shapiro and Nemirovski (2005) illustrate that linear decision rules can reduce the complexity of multistage stochastic programming problems. Kuhn et al. (2009) proposes an efficient method to estimate the loss of optimality incurred by linear decision rule approximation.

In this strand of literature, we believe the closest works to the methodology described in our paper are Calafiore (2009) and Skaf and Boyd (2010). Both of these papers use linear decision rules to address dynamic portfolio choice problems with proportional transaction costs without return predictability. Calafiore (2009) compute lower and upper bounds on the expected transaction costs and solves two convex optimization problems to get upper and lower bounds on the optimal value of the simplified dynamic optimization program with linear decision rules. On the other hand, Skaf and Boyd (2010) study the dynamic portfolio choice problem as an application to their general methodology of using affine controllers on convex stochastic programs. They first linearize the dynamics of the wealth process and then solve the resulting convex optimization via sampling techniques. The foremost difference between our approach and these papers is the modeling of return predictability. Hence, the optimal rebalancing rule in our model is a linear function of the predicting factors. Furthermore, we derive exact reductions to deterministic convex programs in the cases of proportional and nonlinear transaction costs.

2. Dynamic Portfolio Choice with Return Predictability and Transaction Costs

We consider a dynamic portfolio choice problem with allowing general models for the predictability of security returns and for trading frictions. The number of investable securities is $N$, time is discrete and indexed by $t = 1, \ldots, T$, where $T$ is the investment horizon. Each security $i$ has a price change $r_{i,t+1}$ from time $t$ to $t+1$.

We collect these price changes in the return vector $r_{t+1} \triangleq (r_{1,t+1}, \ldots, r_{N,t+1})$. We assume that the investor has a predictive model of future security returns, and that these predictions are made through a set of $K$ return-predictive factors. These factors could be security specific characteristics such as the market capitalization of the stock, the book-to-market ratio of the stock, the lagged twelve month return of the stock (see, e.g., Fama and French, 1996; Goetzmann and Jorion, 1993). Alternatively, they could be macroeconomic signals that affect the return of each security,  

\footnote{We choose to describe the evolution of asset prices in our framework in terms of absolute price changes, and we will also refer to these as (absolute) returns. Note that this is without loss of generality: since the return dynamics specified by Assumption \ref{assumption} allow for an arbitrary dependence on history, our framework also admits, for example, models which describe the rate of return of each security.}
such as inflation, treasury bill rate, industrial production (see, e.g., Chen et al., 1986). Denote by $f_t \in \mathbb{R}^K$ the vector of factor values at time $t$. We assume very general dynamics, possibly nonlinear and with a general dependence on history, for the evolution of returns and factors.

**Assumption 1 (General return and factor dynamics).** Over a complete filtered probability space given by $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \geq 0}, P)$, we assume that factors and returns evolve according to

$$f_{t+1} = G_{t+1}(f_t, \ldots, f_1, \epsilon_{t+1}), \quad r_{t+1} = H_{t+1}(f_t, \epsilon_{t+1}),$$

for each time $t$. Here, $G_{t+1}(\cdot)$ and $H_{t+1}(\cdot)$ are known functions that describe the evolution of the factors and returns in terms of the history of factor values and the exogenous, i.i.d. disturbances $\epsilon_{t+1}$. We assume that the filtration $\mathcal{F} \triangleq \{\mathcal{F}_t\}_{t \geq 0}$ is the natural filtration generated by the exogenous noise terms $\{\epsilon_t\}$.

Let $x_{i,t}$ denote the number of shares that the investor holds in $i$th security over the time period $t$. We collect the portfolio holdings across all securities at time $t$ in the vector $x_t \triangleq (x_{1,t}, \ldots, x_{N,t})$, and we denote the fixed initial portfolio of the investor by $x_0$. Similarly, let the trade vector $u_t \triangleq (u_{1,t}, \ldots, u_{N,t})$ denote the amount of shares that the investor wants to trade at the beginning of the $t$th period, when he inherits the portfolio $x_{t-1}$ from the prior period and observes the latest realization of factor values $f_t$. Consequently, we have the following linear dynamics for our position and trade vector:

$$x_t = x_{t-1} + u_t,$$

for each $t$.

Let the entire sample path of portfolio positions, factor realizations, and security returns be denoted by $x \triangleq (x_1, \ldots, x_T), \triangleq (f_1, \ldots, f_T)$, and $r \triangleq (r_2, \ldots, r_{T+1})$, respectively. Similarly, the sample path of trades over time is denoted by $u = (u_1, \ldots, u_T)$. We make the following assumption on feasible sample paths of trades:

**Assumption 2 (Convex trading constraints).** The sample path of trades $u$ are restricted to the non-empty, closed, and convex set $U \subseteq \mathbb{R}^N \times \ldots \times \mathbb{R}^N$.

The investor’s trading decisions are determined by a policy $\pi$ that selects a sample path of trades $u$ in $U$ for each realization of $r$ and $f$. We let $\mathcal{U}$ be the set of all policies. We assume that the investor’s trading decisions are non-anticipating in that the trade vector $u_t$ in period $t$ depends only on what is known at the beginning of period $t$. Formally, we require policies to be adapted to the filtration $\mathcal{F}$, such that a policy’s selection of the trade vector $u_t$ at time $t$ must be measurable with respect to $\mathcal{F}_t$. Let $\mathcal{U}_F$ be the set of all non-anticipating policies.

The objective of the investor is to select a policy $\pi \in \mathcal{U}_F$ that maximizes the expected value of a total reward or payoff function $p(\cdot)$. Formally, we consider the following optimization problem for the investor,

$$(1) \quad \sup_{\pi \in \mathcal{U}_F} \mathbb{E}_{\pi}[p(x, f, r)],$$

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where the real-valued reward function \( p(\cdot) \) is a function of the entire sample path of portfolio positions, \( x \), the factor realization, \( f \), and security returns \( r \). For example, \( p(\cdot) \) may have the form

\[
p(x, f, r) \triangleq W(x, r) - TC(u) - RA(x, f, r).
\]

Here, \( W \) denotes the terminal wealth (total trading gains ignoring of transaction costs), i.e.,

\[
W(x, r) \triangleq W_0 + \sum_{t=1}^{T} x_t^\top r_{t+1},
\]

where \( W_0 \) is the initial wealth. \( TC(\cdot) \) captures the transaction costs associated with a set of trading decisions, and \( RA(\cdot) \) is the penalty term that incorporates risk aversion.

We make the following assumption about our objective function:

**Assumption 3 (Concave objective function).** *Given an arbitrary, fixed sample paths of factor realizations \( f \) and security returns \( r \), assume that the reward function \( p(x, f, r) \) is a concave function of the sequence of positions \( x \).*

If \( p(\cdot) \) has the specified form in (2), then Assumption 3 will be satisfied when the transaction cost term \( TC(\cdot) \) is a convex function of trades and the risk aversion term \( RA(\cdot) \) is a convex function of portfolio positions.

### 2.1. Examples

In this paper, we consider dynamic portfolio choice models that satisfy Assumptions 1–3. In order to illustrate the generality of this setting, we will now provide a number of specific examples that satisfy these assumptions.

**Example 1** ([Garleanu and Pedersen 2009]). *This model has the following dynamics, where returns are driven by mean-reverting factors, that fit into our general framework:

\[
\begin{align*}
    f_{t+1} &= (I - \Phi) f_t + \epsilon^{(1)}_{t+1}, \\
    r_{t+1} &= \mu_t + B f_t + \epsilon^{(2)}_{t+1},
\end{align*}
\]

for each time \( t \geq 0 \). Here, \( \mu_t \) is the deterministic ‘fair return’, e.g., derived from the CAPM, while \( B \) is a matrix of constant factor loadings. The factor process \( f_t \) is a vector mean-reverting process, with \( \Phi \) a matrix of mean reversion coefficients for the factors. It is assumed that the i.i.d. disturbances \( \epsilon_{t+1} \triangleq (\epsilon^{(1)}_{t+1}, \epsilon^{(2)}_{t+1}) \) are zero-mean with covariance given by \( \text{Var}(\epsilon^{(1)}_{t+1}) = \Psi \) and \( \text{Var}(\epsilon^{(2)}_{t+1}) = \Sigma \).

Trading is costly, and the transaction cost to execute \( u_t = x_t - x_{t-1} \) shares is given by \( TC_t(u_t) \triangleq 1/2 u_t \Lambda u_t \), where \( \Lambda \in \mathbb{R}^{N \times N} \) is a positive semi-definite matrix that measures the level of trading costs. There are no trading constraints (i.e., \( U \triangleq \mathbb{R}^{N \times T} \)). The investor’s objective function is to choose a trading strategy to maximize discounted future expected excess return, while accounting for
transaction costs and adding a per-period penalty for risk, i.e.,

\begin{equation}
\maximize_{\pi \in \mathcal{U}} \mathbb{E}[\pi] \left[ \sum_{t=1}^{T} \left( x_t^\top B f_t - TC_t(u_t) - RA_t(x_t) \right) \right].
\end{equation}

where \( RA_t(x_t) \triangleq \frac{\gamma^2}{2} x_t^\top \Sigma x_t \) is a per-period risk aversion penalty, with \( \gamma \) being a coefficient of risk aversion. Garleanu and Pedersen (2009) suggest this objective function for an investor who is compensated based on his performance relative to a benchmark. Each \( x_t^\top B f_t \) term measures the excess return over the benchmark, while each \( RA_t(x_t) \) term measures the variance of the tracking error relative to the benchmark.\(^2\)

The problem (4) clearly falls into our framework. The objective function is similar to that of (2) with the minor variation expected excess return rather than expected wealth is considered. Further, (4) has the further special property that total transaction costs and penalty for risk aversion decompose over time:

\[ RA(x, f, r) \triangleq \sum_{t=1}^{N} RA_t(x_t), \quad TC(u) \triangleq \sum_{t=1}^{N} TC_t(u_t). \]

Note that this problem can be handled easily using the classical theory from the linear-quadratic control (LQC) literature (see, e.g., Bertsekas, 2000). This theory provides analytical characterization of optimal solution, for example, that the value function at any time \( t \) is quadratic function the state \( (x_t, f_t) \), and that the optimal trade at each time is an affine function of the state. Moreover, efficient computational procedures are available to solve for the optimal policy.

On the other hand, the tractability of this model rests critically on three key requirements:

- The state variables \( (x_t, f_t) \) at each time \( t \) must evolve as linear functions of the control \( u_t \) and the i.i.d. disturbances \( \epsilon_t \) (i.e., linear dynamics).
- Each control decision \( u_t \) is unconstrained.
- The objective function must decompose across time into a positive definite quadratic function of \( (x_t, u_t) \) at each time \( t \).

These requirements are not satisfied by many real world examples, which may involve portfolio position or trade constraints, different forms of transaction costs and risk measures, and more complicated return dynamics. In the following examples, we will provide concrete examples of many such cases that do not admit optimal solutions via the LQC methodology, but remain within our framework.

**Example 2 (Portfolio or trade constraints).** In practice, a common constraint in constructing equity portfolios is the short-sale restriction. Most of the mutual funds are enforced not to have any short

\(^2\)See Garleanu and Pedersen (2009) for other interpretations.
positions by law. This requires the portfolio optimization problem to include the linear constraint

$$x_t = x_0 + \sum_{s=1}^{t} u_t \geq 0,$$

for each $t$. This is clearly a convex constraint on the set of feasible trade sequence $u$.

We observe a similar restriction when an execution desk needs to sell or buy a large portfolio on behalf of an investor. Due to the regulatory rules in agency trading, the execution desk is only allowed to sell or buy during the trading horizon. In the ‘pure-sell’ scenario, the execution desk needs to impose the negativity constraint

$$u_t \leq 0,$$

for each time $t$.

Simple linear constraints such as these fit easily in our framework, but cannot be addressed via traditional LQC methods.

**Example 3 (Non-quadratic transaction costs).** In practice, many trading costs such as the bid-ask spread, broker commissions, and exchange fees are intrinsically proportional to the trade size. Letting $\chi_i$ be the the proportional transaction cost rate (an aggregate sum of bid-ask cost and commission fees, for example) for trading security $i$, the investor will incur a total cost of

$$TC(u) \triangleq \sum_{t=1}^{T} \sum_{i=1}^{N} \chi_i |u_{i,t}|.$$

The proportional transaction costs are a classical cost structure that is well studied in the literature (see, e.g., Constantinides, 1986).

Furthermore, other trading costs occur due to disadvantageous transaction price caused by the price impact of the trade. The management of the trading costs due to price impact has recently attracted considerable interest (see, e.g., Obizhaeva and Wang, 2005; Almgren and Chriss, 2000). Many models of transaction costs due to price impact imply a nonlinear relationship between trade size and the resulting transaction cost, for example

$$TC(u) \triangleq \sum_{t=1}^{T} \sum_{i=1}^{N} \chi_i |u_{i,t}|^\beta.$$

Here, $\beta \geq 1^3$ and $\chi_i$ is a security specific proportionality constant.

In general, when the trade size is small relative to the total traded volume, proportional costs will dominate. On the other hand, when the trade size is large, costs due to price impact will dominate. Hence, both of these types of trading are important. However, the LQC framework of Example 1 only allows quadratic transaction costs (i.e., $\beta = 2$).

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Footnote: $^3$Gatheral (2010) notes that $\beta = \frac{3}{2}$ is a typical assumption in practice.
Example 4 (Terminal wealth risk). The objective function of Example [1] includes a term to penalize excessive risk. In particular, the per-period quadratic penalty, $x_t^\top \Sigma x_t$, is used, in order to satisfy the requirements of the LQC model. However, penalizing additively risk in a per-period fashion is nonstandard. Such a risk penalty does not correspond to traditional forms of investor risk preferences, e.g., maximizing the utility of terminal wealth, and the economic meaning of such a penalty is not clear. An investor is typically more interested in the risk associated with the terminal wealth, rather than a sum of per-period penalties.

In order to account for terminal wealth risk, let $\rho: \mathbb{R} \to \mathbb{R}$ be a real-valued convex function meant to penalize for excessive risk of terminal wealth (e.g., $\rho(w) = \frac{1}{2}w^2$ for a quadratic penalty) and consider the optimization problem

$$
\maximize_{\pi \in \mathcal{U}_\pi} \mathbb{E}_\pi \left[ W(x, r) - TC(u) - \gamma \rho(W(x, r)) \right],
$$

where $\gamma > 0$ is a risk-proportionality constant.

It is not difficult to see that the objective in (5) satisfies Assumption 3 and hence fits into our model. However, even when the risk penalty function $\rho(\cdot)$ is quadratic, (5) does not admit a tractable LQC solution, since the quadratic objective does not decompose across time.

Example 5 (Maximum drawdown risk). In addition to the terminal measures of risk described in Example 4, an investor might also be interested controlling intertemporal measures of risk defined over the entire time trajectory. For example, a fund manager might be sensitive to a string of successive losses that may lead to the withdrawal of assets under management. One way to limit such losses is to control the maximum drawdown, defined as the worst loss of the portfolio between any two points of time during the investment horizon $^4$. Formally,

$$
\text{MD}(x, r) \triangleq \max_{1 \leq t_1 \leq t_2 \leq T} \left( - \sum_{t=t_1}^{t_2} x_t^\top r_{t+1}, 0 \right).
$$

It is easy to see that the maximum drawdown is a convex function of $x$. Hence, the portfolio optimization problem

$$
\maximize_{\pi \in \mathcal{U}_\pi} \mathbb{E}_\pi \left[ W(x, r) - TC(u) - \gamma \text{MD}(x, r) \right],
$$

where $\gamma \geq 0$ is a constant controlling tradeoff between wealth and the maximum drawdown penalty, satisfies Assumption 3. Moreover, standard convex optimization theory yields that the problem (6) is equivalent to solving the constrained problem

$$
\maximize_{\pi \in \mathcal{U}_\pi} \mathbb{E}_\pi \left[ W(x, r) - TC(u) \right] \quad \text{subject to} \quad \mathbb{E}_\pi \left[ \text{MD}(x, r) \right] \leq C,
$$

$^4$For example, see Grossman and Zhou (1993) for an earlier example.
where $C$ (which depends on the choice of $\gamma$) is a limit on the allowed expected maximum drawdown.

**Example 6 (Complex dynamics).** We can also generalize the dynamics of Example 1. Consider factor and return dynamics given by

$$f_{t+1} = (I - \Phi) f_t + \epsilon_{t+1}^{(1)}, \quad r_{t+1} = \mu_t + (B + \xi_{t+1}) f_t + \epsilon_{t+1}^{(2)},$$

for each time $t \geq 0$. Here, each $\xi_{t+1} \in \mathbb{R}^{N \times K}$ is an extra noise term which captures model uncertainty regarding the factor loadings. We assume that

$$\mathbb{E}[(B + \xi_{t+1}) f_t \mid \mathcal{F}_t] = B f_t, \quad \text{Var}[(B + \xi_{t+1}) f_t \mid \mathcal{F}_t] = f_t^\top \Upsilon f_t.$$

With this model, the conditional variance of returns becomes dependent on the factor structure and is time-varying, i.e., $\text{Var}[r_{t+1} \mid \mathcal{F}_t] = f_t^\top \Upsilon f_t + \Sigma$. This is consistent with the empirical work of Fama and French (1996), for example. In this setting, the per-period penalty of risk analogous to that in (1) becomes $RAR_t(x, f) = x_t^\top \left(f_t^\top \Upsilon f_t + \Sigma\right) x_t$. The resulting optimal control problem no longer falls into the LQC framework.

The dynamics and the reward functions considered in these examples satisfy our basic requirements of Assumptions 1–3. These examples illustrate that in many real-world problems with complex primitives for return predictability, transaction costs, risk measures and constraints, the dynamic portfolio choice becomes difficult to solve analytically via LQC methods.

### 3. Optimal Linear Model

The examples of Section 2.1 illustrated a broad range of practically important portfolio optimization problems. Without special restrictions, such as those imposed in the LQC framework, the optimal dynamic policy for such a broad set of problems cannot be computed either analytically or computationally. In this section, in order to obtain policies in a computationally tractable way, we will consider a more modest goal. Instead of finding the optimal policy amongst all admissible dynamic policies, we will restrict our search to a subset of policies that are parsimoniously parameterized. That is, instead of solving for a globally optimal policy, we will instead find an approximately optimal policy by finding the best policy over the restricted subset of policies.

In order to simplify, we will assume that reward function of the investor’s optimization (1) is a function only of the sample path of portfolio positions $x$ and of factor realizations $f$, and does not depend on the security returns $r$. In other words, we assume that the reward function takes the form $p(x, f)$. This is without loss of generality — given our general specification for factors under Assumption 1, we can simply include each security return as a factor. With this assumption, investor’s trading decisions will, in general, be a non-anticipating function of the sample path of factor realizations $f$. However, consider the following restricted set of policies, *linear rebalancing policies*, which are obtained by taking the affine combinations of the factors:
**Definition 1 (Linear rebalancing policy).** A linear rebalancing policy $\pi$ is a non-anticipating policy parameterized by collection of vectors $c \triangleq \{c_t \in \mathbb{R}^N, 1 \leq t \leq T\}$ and a collection of matrices $E \triangleq \{E_{s,t} \in \mathbb{R}^{N \times K}, 1 \leq s \leq t \leq T\}$, that generates a sample path of trades $u \triangleq (u_1, \ldots, u_T)$ according to

$$u_t \triangleq c_t + \sum_{s=1}^{t} E_{s,t} f_s,$$

for each time $t = 1, 2, \ldots, T$.

Define $C$ to be the set of parameters $(E, c)$ such that the resulting sequence of trades $u$ is contained in the constraint set $U$, with probability 1, i.e., $u$ is feasible. Denote by $\mathcal{L} \subset \mathcal{U}_\mathcal{F}$ the corresponding set of feasible linear policies.

Observe that linear rebalancing allow recourse, albeit in a restricted functional form. The affine specification (8) includes several classes of polices of particular interest as special cases:

- **Deterministic policies.** By taking $E_{s,t} \triangleq 0$, for all $1 \leq s \leq t \leq T$, it is easy to see that any deterministic policy is a linear rebalancing policy.

- **LQC optimal policies.** Optimal portfolios for the LQC framework of Example 1 take the form

$$x_t = \Gamma_{x,t} x_{t-1} + \Gamma_{f,t} f_t,$$

given matrices $\Gamma_{x,t} \in \mathbb{R}^{N \times N}$, $\Gamma_{f,t} \in \mathbb{R}^{N \times K}$, for all $1 \leq t \leq T$, i.e., the optimal portfolio are linear in the previous position and the current factor values. Equivalently, by induction on $t$,

$$x_t = \left( \prod_{s=1}^{t} \Gamma_{x,s} \right) x_0 + \sum_{s=1}^{t} \left( \prod_{\ell=1}^{s-1} \Gamma_{x,\ell} \right) \Gamma_{f,s} f_s.$$

Since $u_t = x_t - x_{t-1}$, it is clear that the optimal trade $u_t$ is a linear function of the fixed initial position $x_0$, and the factor realizations $\{f_1, \ldots, f_t\}$, and is therefore of the form (8).

- **Linear portfolio polices.** Brandt et al. (2009) suggest a class of policies where portfolios are determined by adjusting a deterministic benchmark portfolio according to a linear function of a vector of stochastic, time-varying firm characteristics. In our setting, the firm characteristics would be interpreted as stochastic return predicting factors. An analogous rule would determine the positions at each time $t$ via

$$x_t = \bar{x}_t + \Theta_{t}^\top (f_t - \bar{f}_t).$$

Here, $\bar{f}_t$ is the expected factor realization at time $t$. The policy is parameterized by $\bar{x}_t$, the deterministic benchmark portfolio at time $t$, and the matrix $\Theta_t \in \mathbb{R}^{N \times K}$, which maps firm characteristics (standardized to be mean zero) to adjustments to the benchmark portfolio. Such a portfolio rule is clearly of the form (8).
• **Policies based on basis functions.** Instead of having policies that are directly affine function of factor realizations, it is also possible to introduce basis functions. One might consider, for example, \( \varphi: \mathbb{R}^K \rightarrow \mathbb{R}^D \), a collection of \( D \) (non-linear) functions that capture particular features of the factor space that are important for good decision making. Consider a class of policies of the form

\[
u_t \triangleq c_t + \sum_{s=1}^{t} E_{s,t} \varphi(f_s).
\]

Such policies belong to the linear rebalancing class, if the factors are augmented also to include the value of the basis functions. This is easily accommodated in our framework, given the flexibility of Assumption 1. Similarly, policies which depend on the past security returns (in addition to factor realizations) can be accomodated by augmenting the factors with past returns.

An alternative to solving the original optimal control problem (1) is to consider the problem

\[
\sup_{\pi \in \mathcal{L}} \mathbb{E}[p(x, f)],
\]

which restricts to linear rebalancing rules. In general, (9) will not yield an optimal control. The exception is if the optimal control for the problem is indeed a linear rebalancing rule (e.g., in a LQC problem). However, (9) will yield the best possible linear rebalancing rule. Further, in contrast to the original optimal control problem, (9) has the great advantage of being *tractable*, as suggested by the following result:

**Proposition 1.** The optimization problem given by

\[
\begin{align*}
\text{maximize} & \quad \mathbb{E}[p(x, f)] \\
\text{subject to} & \quad x_t = x_{t-1} + u_t, \quad \forall \ 1 \leq t \leq T, \\
& \quad u_t = c_t + \sum_{s=1}^{t} E_{s,t} f_s, \quad \forall \ 1 \leq t \leq T, \\
& \quad (E, c) \in \mathcal{C}.
\end{align*}
\]

is a convex optimization problem, i.e., it involves the maximization of a concave function subject to convex constraints.

**Proof.** Note that \( p(\cdot, f) \) is concave for a fixed \( f \) by Assumption 3 and since \( x \) can be written as an affine transformation of \((E, c)\). Then, for each fixed \( f \), the objective function is concave in \((E, c)\). Taking an expectation over \( f \) preserves this concavity. Finally, the convexity of the constraint set \( \mathcal{C} \) follows from the convexity of \( U \), under Assumption 2. 

The problem (10) is a finite-dimensional, convex optimization problem that will yield parameters for the optimal linear rebalancing policy. It is also a stochastic optimization problem, in the sense
that the objective is the expectation of a random quantity. In general, there are a number of effective numerical methods that can been applied to solve such problems:

- **Efficient exact formulation.** In many cases, with further assumptions on the problem primitives (the reward function \( p(\cdot) \), the dynamics of the factor realizations \( f \), and the trading constraint set \( U \)), the objective \( E[p(x, f)] \) and the constraint set \( C \) of the program (10) can be explicitly analytically expressed in terms of the decision variables \((E, c)\). In some of these cases, the program (10) can be transformed into a standard form of convex optimization program such as a quadratic program or a second-order cone program. In such cases, off-the-shelf solvers specialized to these standard forms (e.g., Grant and Boyd, 2011) can be used. Alternatively, generic methods for constrained convex optimization such as interior point methods (see, e.g., Boyd and Vandenberghe, 2004) can be applied to efficiently solve large-scale instances of (10). We will explore this topic further, developing a number of efficient exact formulations in Section 4 and providing a numerical example in Section 5.

- **Sample average approximation (SAA).** In the absence of further structure on the problem primitives, the program (10) can also be solved via Monte Carlo sampling. Specifically, supposed that \( f^{(1)}, \ldots, f^{(S)} \) are \( S \) independent sample paths of factor realization. The objective and constraints of (10) can be replaced with sampled versions, to obtain

\[
\begin{align*}
\text{maximize} & \quad \frac{1}{S} \sum_{\ell=1}^{S} p(x^{(\ell)}, f^{(\ell)}) \\
\text{subject to} & \quad x_t^{(\ell)} = x_{t-1}^{(\ell)} + u_t^{(\ell)}, \quad \forall 1 \leq t \leq T, 1 \leq \ell \leq S, \\
& \quad u_t^{(\ell)} = c_t + \sum_{s=1}^{t} E_{s,t} f_s^{(\ell)}, \quad \forall 1 \leq t \leq T, 1 \leq \ell \leq S, \\
& \quad u^{(\ell)} \in U, \quad \forall 1 \leq \ell \leq S.
\end{align*}
\]

The sample average approximation (11) can be solved via standard convex optimization methods (e.g., interior point methods). Moreover, under appropriate regularity conditions, convergence of the SAA (11) to the original program (10) can be established as \( S \to \infty \), along with guarantees on the rate of convergence (Shapiro, 2003).

- **Stochastic approximation.** Denote the collection of decision variables in (10) by \( z \triangleq (E, c) \), and, allowing a minor abuse of notation, define \( p(z, f) \) to be the reward when the sample path of factor realizations is given by \( f \) and the trading policy is determined by \( z \). Then, defining \( h(z) \triangleq p(z, f) \), the problem (10) is simply to maximize \( h(z) \) subject to the constraint that \( z \in C \). Under suitable technical conditions, superdifferentials of \( h \) and \( p \) are related according to

\[ \partial h(z) = E[\partial_z p(z, f)]. \]

Stochastic approximation methods are incremental methods that seek to estimate ascent directions for \( h(\cdot) \) from sampled ascent directions for \( p(\cdot, f) \). For example, given a sequence
of i.i.d. sample paths of factor realizations $\mathbf{f}^{(1)}, \mathbf{f}^{(2)}, \ldots$, a sequence of parameter estimates $\mathbf{z}^{(1)}, \mathbf{z}^{(2)}, \ldots$ can be constructed according to

$$
\mathbf{z}^{(\ell+1)} = \Pi_C \left( \mathbf{z}^{(\ell)} + \gamma \ell \zeta_\ell \right),
$$

where $\Pi_C(\cdot)$ is the projection onto the feasible set $C$, $\zeta_\ell \in \partial \mathcal{p}(\mathbf{z}^{(\ell)}; \mathbf{f}^{(\ell)})$ is a supergradient, and $\gamma_\ell > 0$ is a step-size. Stochastic approximation methods have the advantage of being incremental and thus requiring minimal memory relative to sample average approximation, and are routinely applied in large scale convex stochastic optimization [Nemirovski et al., 2009].

One attractive feature of the sample average approximation and stochastic approximation approaches is that they can be applied in a data-driven fashion. These methods need access only to simulated trajectories of factors and returns — they do not need explicit knowledge of the dynamics in Assumption 1 that drive these processes. Hence, an optimal linear rebalancing policy can be determined using, for example, historical data to construct simulated trajectories, without specifying and estimating an explicit functional form for the factor and return dynamics.

Finally, observe that optimal linear policies can also be applied in concert with model predictive control (MPC). Here, at each time step $t$, the program (10) is resolved beginning from time $t$. This determines the optimal linear rebalancing rule from time $t$ forward, conditioned on the realized history up to time $t$. The resulting policy is only used to determine the trading decision at the then current time $t$, and (10) is subsequently resolved at each future time period. At the cost of an additional computational burden, the use of optimal linear policies with MPC subsumes standard MPC approaches, such as resolving a myopic variation of the portfolio optimization problem (and ignoring the true multiperiod nature) or solving a deterministic variation of the portfolio optimization problem (and ignoring the possibility of future recourse).

4. Efficient Exact Formulations

In this section, we will provide efficient exact formulations of dynamic portfolio choice problems using the class of linear policies for our feasible set of policies. In particular, we will consider a number of the examples of dynamic portfolio choice problems discussed in Section 2.1. These examples include features such as constraints on portfolio holdings, transaction costs, and risk measures. In each case, we will demonstrate how the optimization problem (10) can be transformed into a deterministic convex program by explicit analytical evaluation of the objective function $\mathbb{E}[p(\cdot; \mathbf{f})]$ and the constraint set $C$.

Exact formulations require the evaluation of expectations taken over the sample path of factor realizations $\mathbf{f}$. In order to do this, we will make the following assumption for the rest of this section:

**Assumption 4 (Gaussian factors).** Assume that the sample path $\mathbf{f}$ of factor realizations is jointly Gaussian. In particular, denote by $\mathbf{F}_t \triangleq (f_1, \ldots, f_t)^\top \in \mathbb{R}^{Kt}$ the vector of all factors observed by
time \( t \). We assume that

\[ F_t \sim N(\theta_t, \Omega_t), \]

where \( \theta_t \in \mathbb{R}^{K_t} \) is the mean vector and \( \Omega_t \in \mathbb{R}^{K_t \times K_t} \) is the covariance matrix.

With this assumption, the trades of any linear policy will also be jointly normally distributed, as each such policy is affine transformations of the factors. Formally, let

\begin{equation}
M_t \triangleq \begin{bmatrix} E_{1,t} & E_{2,t} & \ldots & E_{t,t} \end{bmatrix} \in \mathbb{R}^{N \times K_t}
\end{equation}

be the matrix of time \( t \) policy coefficients, so that the trade vector is given by \( u_t = c_t + M_t F_t \).

With this representation, it is easy see that \( u_t \sim N(\bar{u}_t, V_t) \), where the mean vector and covariance matrix are given by

\begin{equation}
\bar{u}_t \triangleq \mathbb{E}[u_t] = c_t + M_t \theta_t, \quad V_t \triangleq \text{Var}(u_t) = M_t \Omega_t M_t^\top.
\end{equation}

Similarly, the portfolio \( x_t \) at time \( t \) is normally distributed. we have that

\begin{equation}
x_t = x_0 + \sum_{i=1}^{t} u_i = x_0 + \sum_{i=1}^{t} \left( c_i + \sum_{s=1}^{i} E_{s,i} f_s \right) = d_t + \sum_{s=1}^{t} J_{s,t} f_s,
\end{equation}

where \( d_t \triangleq x_0 + \sum_{i=1}^{t} c_i \) and \( J_{s,t} \triangleq \sum_{s=1}^{t} E_{s,t} \). With this representation, it is easy see that \( x_t \sim N(\kappa_t, Y_t) \), where

\begin{equation}
\kappa_t \triangleq \mathbb{E}[x_t] = d_t + P_t \theta_t, \quad Y_t \triangleq \text{Var}(x_t) = P_t \Omega_t P_t^\top,
\end{equation}

and

\begin{equation}
P_t \triangleq \begin{bmatrix} J_{1,t} & J_{2,t} & \ldots & J_{t,t} \end{bmatrix}.
\end{equation}

### 4.1. Linear Constraints

We will provide formulations for linear equality or inequality constraints on trades or positions, in the context of linear rebalancing policies. These type of constraint appear frequently in portfolio choice due to regulatory reasons such as short sale restriction, liquidation purposes or diversification needs such as keeping a specific industry exposure under a certain limit.

#### 4.1.1. Equality Constraints

Equality constraints appear often in portfolio choice, particularly in portfolio execution problems when the investor needs to liquidate a certain portfolio (i.e., \( x_T = 0 \)) or construct a certain target portfolio by the end of the time horizon (i.e., \( x_T = \bar{x} \)).

Suppose that for some time \( t \), have a linear equality constraint on the trade vector \( u_t \), of the
form $Au_t = b$. Here, $A \in \mathbb{R}^{M \times N}$ and $b \in \mathbb{R}^N$. This constraint can be written as

\[(17) \quad Ac_t + AM_tF_t = b.\]

Under Assumption 4, the left hand side of the (17) is normally distributed. Therefore, for (17) to hold almost surely, we must have that the left hand side have mean $b$ and zero covariance. Thus, we require that

\[(18) \quad Ac_t = b, \quad AM_t = 0.\]

Thus, the linear equality constraint (17) on the trade vector $u_t$ is equivalent to the linear equality constraint (18) on the policy coefficients $(c_t, M_t)$. Linear equality constraints on the portfolio position $x_t$ can be handled similarly.

### 4.1.2. Inequality Constraints

Inequality constraints on trades or positions are common as well. One example is a short-sale constraint, which would require that $x_t \geq 0$ for all times $t$. When the factor realizations do not have bounded support, inequality constraints cannot be enforced almost surely. This is true in the Gaussian case: there is a chance, however small, that factors may take extreme values, and if the policy if a linear function of the factors, this may cause an inequality constraint to be violated.

In order to account for such constraints in a linear rebalancing policy, instead of enforcing inequality constraints almost surely, we will enforce them at a given level of confidence. For example, given a vector $a \in \mathbb{R}^N$ and a scalar $b$, instead of enforcing the linear constraint

$$a^T u_t \leq b,$$

almost surely, we can consider a relaxation where seek to guarantee that it is violated with small probability. In other words, we can impose the chance constraint

$$\mathbb{P}(a^T u_t > b) \leq \eta,$$

for a small value of the parameter $\eta$. The following lemma, whose proof can be found in Appendix A, illustrates that this can be accomplished explicitly:

**Lemma 1.** Given $\eta \in [0, 1/2]$, a non-zero vector $a \in \mathbb{R}^N$, and a scalar $b$, the chance constraint

$$\mathbb{P}(a^T u_t > b) \leq \eta$$

is equivalent to the constraint

$$a^T (c_t + M_t \theta_t) - b + \Phi^{-1}(1 - \eta) \left\| \Omega_t^{1/2} M_t^T a \right\|_2 \leq 0$$

on the policy coefficients $(c_t, M_t)$, where $\Phi^{-1}(\cdot)$ is the inverse cumulative normal distribution.

A similar approach be applied to incorporate linear inequality constraints on the portfolio
position \( x_t \) with high confidence.

In many situations (e.g., short-sale constraints), it may not be sufficient to enforce an inequality constraint only probabilistically. In such cases, when a linear rebalancing policy is applied, the resulting trades can be projected onto the constraint set so as to ensure that the constraints are always satisfied. When the linear policy is designed, however, it is helpful to incorporate the desired constraints probabilistically so as to account for their presence. We will demonstrate this idea in the application in Section 5.

4.2. Transaction Costs

In this section, we will provide efficient exact formulations for the transaction cost functions discussed in Section 2.1 in the context of linear rebalancing policies. In general, once might consider a total transaction cost of

\[
TC(u) \triangleq \sum_{t=1}^{T} TC_t(u_t)
\]

for executing the sample path of trades \( u \), where \( TC_t(u_t) \) is the cost of executing the trade vector \( u_t \) at time \( t \). As seen in Section 2.1, we typically wish to subtract an expected transaction cost term from investor’s objective. Hence efficient exact formulations for transaction costs involve explicit analytical computation of \( E[TC(u)] = \sum_{t=1}^{T} E[TC_t(u_t)] \), when each trade vector \( u_t \) is specified by a linear policy.

Under a linear policy, \( u_t \sim N(\bar{u}_t, V_t) \) is distributed as a normal random variable, with mean and covariance \( (\bar{u}_t, V_t) \) specified from the policy \((E, c)\) coefficients through (13). Then, the evaluation of expected transaction costs reduces to the evaluation of the expected value of the per period transaction cost function \( TC_t(\cdot) \) for a Gaussian argument. This can be handled on a case-by-case basis as follows:

- **Quadratic transaction costs.** In the case of quadratic transaction costs, as seen in Example 1, the per period transaction cost function is given by

\[
TC_t(u_t) \triangleq \frac{1}{2} u_t^\top \Lambda u_t,
\]

where \( \Lambda \in \mathbb{R}^{N \times N} \) is a positive definite matrix. In this case,

\[
E[TC_t(u_t)] = \frac{1}{2} (\bar{u}_t^\top \Lambda \bar{u}_t + \text{tr}(\Lambda V_t)).
\]

- **Proportional transaction costs.** In the case of proportional transaction costs, as discussed in Example 3, the per period transaction cost function is given by

\[
TC_t(u_t) \triangleq \sum_{i=1}^{N} \chi_i |u_{t,i}|,
\]

where \( \chi_i > 0 \) is a proportionality constant specific to security \( i \). Using the properties of the
folded normal distribution, we obtain

\[
E[T_{C,t}(u_t)] = \sum_{i=1}^{N} \chi_i \left( \sqrt{\frac{2V_{t,i}}{\pi}} \exp \left\{ -\frac{\bar{u}_{t,i}^2}{2V_{t,i}} \right\} + \bar{u}_{t,i} \left\{ 1 - 2\Phi \left( -\frac{\bar{u}_{t,i}}{\sqrt{V_{t,i}}} \right) \right\} \right),
\]

where \( \Phi(\cdot) \) is the cumulative distribution function of a standard normal random variable.

- **Nonlinear transaction costs.** In the case of nonlinear transaction costs, as discussed in Example 3, the per period transaction cost function is given by

\[
T_{C,t}(u_t) \triangleq \sum_{i=1}^{N} \chi_i |u_{t,i}|^\beta,
\]

where \( \chi_i > 0 \) is a proportionality constant specific to security \( i \), and \( \beta \geq 1 \) is an exponent capturing the degree of nonlinearity. As in the proportional case, evaluating the Gaussian expectation explicitly results in

\[
E[T_{C,t}(u_t)] = \sum_{i=1}^{N} \chi_i \Gamma \left( \frac{1 + \beta}{2} \right) \frac{(2V_{t,i})^{\frac{\beta}{2}}}{\sqrt{\pi}} \text{I}_1 \left( -\frac{\beta}{2} ; \frac{1}{2} ; -\frac{\bar{u}_{t,i}^2}{2V_{t,i}} \right),
\]

where \( \Gamma(\cdot) \) is the gamma function and \( \text{I}_1(\cdot) \) is the confluent hypergeometric function of the first kind.

### 4.3. Terminal Wealth and Risk Aversion

In many of the portfolio choice examples in Section 2.1, an investor wishes to maximize expected wealth net of transaction costs, subject to a penalty for risk, i.e.,

\[
\begin{align*}
\text{maximize} & \quad E_{\pi} \left[ W(x, r) - T_{C}(u) - R_{A}(x, f, r) \right].
\end{align*}
\]

Here, \( W(\cdot) \) is the terminal wealth associated with a sample path, \( T_{C}(\cdot) \) are the transaction costs, and \( R_{A}(\cdot) \) is a penalty for risk aversions. Exact calculation of expected transaction costs for linear policies were discussed in Section 4.2. Here, we will discuss exact calculation of the expected terminal wealth and the risk aversion penalty.

To begin, note that the terminal wealth depends on realized returns in addition to factor realizations. Hence, we will make the following assumption:

**Assumption 5 (Gaussian returns).** As in Example 7, assume that for each time \( t \geq 0 \), returns evolve according to

\[
r_{t+1} = \mu_t + B f_t + \varepsilon_t^{(2)},
\]

where \( \mu_t \) is a deterministic vector, \( B \) is a matrix of factor loadings, and \( \varepsilon_t^{(2)} \) are zero-mean i.i.d. Gaussian disturbances with covariance \( \Sigma \).
Note that the critical assumption we are making here is that the factor realizations $f$ and the sample path of security returns $r$ are jointly Gaussian. The particular form (20) is chosen out of convenience but is not necessary.

We can calculate the expected terminal wealth as

$$
E[W(x, r)] = W_0 + \sum_{t=1}^{T} E[x_t^T r_{t+1}] = W_0 + \sum_{t=1}^{T} \left( \mu_t^T \kappa_t + E[x_t^T B f_t] \right),
$$

$$
= W_0 + \sum_{t=1}^{T} \left( \mu_t^T \kappa_t + d_t^T B \delta_t + \sum_{s=1}^{t} (\delta_s^T (B(I - \Phi)^{t-s} J_{s,t}^T) \delta_s + tr(B(I - \Phi)^{t-s} J_{s,t}^T \omega_s)) \right),
$$

where $\omega_s$ is the $s$th $K \times K$ diagonal block matrix of $\Omega_t$.

For the risk aversion penalty, we consider two cases:

- **Per period risk penalty.** Consider risk aversion penalties that decompose over time as

$$
RA(x, f, r) = \sum_{t=1}^{N} RA_t(x_t),
$$

where $RA_t(\cdot)$ is a function which penalizes for risk aversion based on the positions held at time $t$. One such case is the quadratic penalty

$$
RA_t(x_t) \triangleq \frac{\gamma}{2} x_t^T \Sigma x_t
$$

of Example [4] where $\gamma > 0$ is a risk penalty proportionality constant. Here, the investor seeks to penalize in proportion to the conditional per period variance of the portfolio value. So long as the expectation of $RA_t(\cdot)$ can be calculated for Gaussian arguments, then the overall expected risk aversion penalty can be calculated exactly. This can be accomplished for a variety of functions. For example, quadratic penalties can be handled in a manner analogous to the quadratic transaction costs discussed in Section 4.2.

- **Terminal wealth risk penalty.** Alternatively, as discussed in Example [4], a more natural risk aversion criteria might be to penalize risk as a function of the *terminal* wealth. Specifically, an investor with a quadratic utility function would consider a risk aversion penalty

$$
RA(x, f, r) \triangleq -\frac{\gamma}{2} W(x, r)^2,
$$

where $\gamma > 0$ is a risk penalty proportionality constant. We show in Appendix [4] that $E[W(x, r)^2]$ can be analytically computed and the resulting expression is a quadratic convex function of policy coefficients. Note that handling this quadratic penalty enables to accommodate mean-variance type objectives on the terminal wealth.
5. Application: Equity Agency Trading

In this section, we provide an empirical application to illustrate the implementation and the benefits of the optimal linear policy. As our example, we consider an important problem in equity agency trading. Equity agency trading seeks to address the problem faced by large investors such as pension funds, mutual funds, or hedge funds that need to update the holdings of large portfolios. Here, the investor seeks to minimize the trading costs associated with a large portfolio adjustment. These costs, often labeled ‘execution costs’, consist of commissions, bid-ask spreads, and, most importantly in the case of large trades, price impact from trading. Efficient execution of large trades is accomplished via ‘algorithmic trading’, and requires significant technical expertise and infrastructure. For this reason, large investors utilize algorithmic trading service providers, such as execution desks in investment banks. Such services are often provided on an agency basis, where the execution desk trades on behalf of the client, in exchange for a fee. The responsibility of the execution desk is to find a feasible execution schedule over the client-specified trading horizon while minimizing trading costs and aligning with the risk objectives of the client.

The problem of finding an optimal execution schedule has received a lot of attention in the literature since the initial paper of Bertsimas and Lo (1998). In their model, when price impact is proportional to the number of shares traded, the optimal execution schedule is to trade equal number of shares at each trading time. There are number of papers that extend this model to incorporate the risk of the execution strategy. For example, Almgren and Chriss (2000) derive that risk averse agents need to liquidate their portfolio faster in order to reduce the uncertainty of the execution cost.

The models described above seek mainly to minimize execution costs by accounting for the price impact and supply/demand imbalances caused by the investor’s trading. Complementary to this, an investor may also seek to exploit short-term predictability of stock returns to inform the design of a trade schedule. As such, there is a growing interest to model return predictability in intraday stock returns. Often called ‘short-term alpha models’, some of the predictive models are similar to well-known factor models for the study of long-term stock returns, e.g., the Capital Asset Pricing Model (CAPM), or the Fama-French Three Factor Model. Alternatively, short-term predictions can be developed from microstructure effects, for example the imbalance of orders in an electronic limit order book. Heston et al. (2010) document that systematic trading as described in the examples above and institutional fund flows lead to predictable patterns in intraday returns of common stocks.

We will consider an agency trading optimal execution problem in the presence of short-term predictability. One issue that arises here is that, due to the regulatory rules in agency trading, the execution desk is only allowed to either sell or buy a particular security over the course of the trading horizon, depending on whether the ultimate position adjustment desired for that security is negative or positive. However, given a model for short-term predictability, an optimal trading policy that minimizes execution cost may result in both buy and sell trades for the same security as it seeks to exploit short-term signals. Hence, it is necessary to impose constraints on the sign of
trades, as in Example 2.

If an agency trading execution problem has price and factor dynamics which satisfy Assumption 1 and an objective (including transaction costs, price impact, and risk aversion) that satisfies Assumption 3, then we can compute the best execution schedule in the space of linear execution schedules, i.e., the number of shares to trade at each time is a linear function of the previous return predicting factors. We will consider a particular formulation that involves linear price and factor dynamics and a quadratic objective function (as in Example 1). Note that this example does not highlight the full generality of our framework — more interesting cases would involve non-linear factor dynamics (e.g., microstructure-based order imbalance signals) or a non-quadratic objective (e.g., transaction costs as in Example 3). However, this example is intentionally chosen since, in the absence of the trade sign constraint, the problem can be solved exactly with LQC methods. Hence, are able to compare the optimal linear policy to policies derived from LQC methods applied to the unconstrained problem.

The rest of this section is organized as follows. We present our optimal execution problem formulation in Section 5.1. An exact, analytical solution is not available to this problem, hence, in Section 5.2 we describe several approximate solution techniques, including finding the best linear policy. In order to evaluate the quality of the approximate methods, in Section 5.3 we describe several techniques for computing upper bounds on the performance of any policy for our execution problem. In Section 5.4 we describe the empirical calibration of the parameters of our problem. Finally, in Section 5.5 we present and discuss the numerical results.

5.1. Formulation

We follow the general framework of Section 2. Suppose that $x_0 \in \mathbb{R}^N$ denotes the number of shares in each of $N$ securities that we would like to sell before time $T$. We assume that trades can occur at discrete times, $t = 1, \ldots, T$. We define an execution schedule to be the collection $u \triangleq (u_1, \ldots, u_T)$, where each $u_t \in \mathbb{R}^N$ denotes the number of shares traded at time $t$. Note that a negative (positive) value of $u_{i,t}$ denotes a sell (buy) trade of security $i$ at time $t$. The total position at time $t$ is given by $x_t = x_0 + \sum_{s=1}^{t} u_s$.

The formulation of the agency trading optimal execution problem is as follows:

- **Constraints.** Without loss of generality, we will assume that the initial position is positive, i.e., $x_0 > 0$. The execution schedule must liquidate the entire initial position by the end of the time horizon, thus

\begin{equation}
    x_T = x_0 + \sum_{t=1}^{T} u_t = 0.
\end{equation}

Further, agency trading regulations allow only sell trades, thus

\begin{equation}
    u_t \leq 0, \quad t = 1, \ldots, T.
\end{equation}
Note that any schedule satisfying (21)–(22) will also satisfy

\[ x_t = x_0 + \sum_{s=1}^{t} u_s \geq 0, \quad t = 1, \ldots, T. \]

We denote by \( U^0_F \) the set of non-anticipating policies satisfying (21) almost surely, and by \( U_F \) the set of non-anticipating policies satisfying (21)–(23) almost surely.

- **Return and factor dynamics.** We follow the discrete time linear dynamics of Garleanu and Pedersen (2009) as described in Example 1. We assume that the price change of each security from \( t \) to \( t+1 \) is given by the vector \( r_{t+1} \), and is predicted by \( K \) factors collected in a vector \( f_t \). Furthermore, the evolution of factor realizations follow a mean reverting process. Formally, we have the following dynamics for price changes and factor realizations:

\[
    f_{t+1} = (I - \Phi) f_t + \epsilon_{t+1}^{(1)}, \quad r_{t+1} = \mu + B f_t + \epsilon_{t+1}^{(2)},
\]

where \( B \in \mathbb{R}^{N \times K} \) is a constant matrix of factor loadings, \( \Phi \in \mathbb{R}^{K \times K} \) is a diagonal matrix of mean reversion coefficients for the factors, and \( \mu \in \mathbb{R}^{N} \) is the mean return. We assume that the noise terms are i.i.d., and normally distributed with zero-mean and with covariance matrices given by. \( \text{Var}(\epsilon_{t+1}^{(1)}) = \Psi \in \mathbb{R}^{N \times N} \) and \( \text{Var}(\epsilon_{t+1}^{(2)}) = \Sigma \in \mathbb{R}^{K \times K} \). We discuss the precise choice of return predicting factors and the calibration of the dynamics shortly in Section 5.4.

- **Objective.** We assume that the investor is risk-neutral and seeks to maximize total excess profits after quadratic transaction costs, i.e.,

\[
    \mathcal{V}^* \triangleq \maximize_{\pi \in U_F} E_{\pi} \left[ \sum_{t=1}^{T} \left( x_t^T B f_t - \frac{1}{2} u_T^T \Lambda u_t \right) ^2 \right],
\]

where \( \Lambda \in \mathbb{R}^{N \times N} \) is a matrix parameterizing the quadratic transaction costs.

Note that the problem (24) is a special case of the optimization program in Example 1 with the exception of the constraints (21)–(23).

### 5.2. Approximate Policies

Since an exact, analytical solution is not available, we compare four approximate solution techniques to solve the optimal execution problem in (24):

- **Deterministic.** Instead of allowing for a non-anticipating dynamic policy, where the trade at each time \( t \) is allowed to depend on all events that have occurred before \( t \), we can solve for an optimal static policy, i.e., a deterministic sequence of trades over the entire time horizon that

\footnote{Note that Garleanu and Pedersen (2009) consider an infinite horizon setting, while our setting is finite horizon. Further, Garleanu and Pedersen (2009) solve for dynamic policies in the absence of the constraints (21)–(23).}
is decided at the beginning of the time horizon. Here, observe that at the beginning of the time horizon, the expected future factor vector is given by $E[f_t|f_0] = (I - \Phi)^tf_0$. Therefore, in order to find the optimal deterministic policy, given $f_0$, we maximize the conditional expected value of the stochastic objective in (24) by solving the quadratic program

$$\begin{align*}
\max_u & \quad \sum_{t=1}^{T} \left( x_t^T B (I - \Phi)^t f_0 - \frac{1}{2} u_t^T \Lambda u_t \right) \\
\text{subject to} & \quad u_t = x_t - x_{t-1}, \quad t = 1, \ldots, T, \\
& \quad u_t \leq 0, \quad x_t \geq 0, \quad t = 1, \ldots, T, \\
& \quad x_T = 0,
\end{align*}$$

(25)

to yield a deterministic sequence of trades $u$.

- **Model predictive control.** In this approximation, at each trading time, we solve for the deterministic sequence of trades conditional on the available information and implement only the first trade. Thus, this policy is an immediate extension of the deterministic policy, with the addition of resolving at each trading time. Formally, at time $t$, we solve the quadratic program

$$\begin{align*}
\max_{u_t, \ldots, u_T} & \quad \sum_{s=t}^{T} \left( x_s^T B (I - \Phi)^{s-t} f_t - \frac{1}{2} u_s^T \Lambda u_s \right) \\
\text{subject to} & \quad u_s = x_s - x_{s-1}, \quad s = t, \ldots, T, \\
& \quad u_s \leq 0, \quad x_s \geq 0, \quad s = t, \ldots, T, \\
& \quad x_T = 0,
\end{align*}$$

(26)

If $(u_t^*, \ldots, u_T^*)$ is the optimal solution, then the investor trades $u_t^*$ at time $t$.

- **Projected LQC.** If the inequality constraints (22)–(23) are eliminated, the program would reduce to the classical linear quadratic control problem

$$\begin{align*}
\max_{\pi \in \mathcal{U}_0} & \quad E_{\pi} \left[ \sum_{t=1}^{T} \left( x_t^T B f_t - \frac{1}{2} u_t^T \Lambda u_t \right) \right].
\end{align*}$$

(27)

The optimal dynamic policy for the program in (27) yields the trade

$$u_t = (\Lambda + A_{xx,t+1})^{-1} (\Lambda x_{t-1} + (B + A_{xf,t+1} (I - \Phi)) f_t) - x_{t-1}$$

(28)

at each time $t$ as a function of the previous position $x_{t-1}$ and the current factor values $f_t$. Here, the matrices $A_{xx,t+1}$ and $A_{xf,t+1}$ are derived in the Appendix. The dynamic rule for $u_t$ in (28) of course will not be feasible for the constrained program (24), in general. This, the projected LQC policy seeks a trade decision, $\hat{u}_t$, which is the projection of $u_t$ onto the
constraint set \((22)–(23)\), i.e.,

\[
\hat{u}_{i,t} = \max \{-x_{i,t-1}, \min \{0, u_{i,t}\} \},
\]

for each time \(t < T\) and for each security \(i\).

- **Optimal linear.** As formulated in Definition 1, a linear rebalancing policy specifies trades according to

\[
u_t \triangleq c_t + \sum_{s=1}^{t} E_{s,t} f_s,
\]

for each time \(t = 1, 2, \ldots, T\), given parameters \((E, c)\). Due to the linear relationship between position and trade vectors, we can represent the position vector in the similar form, i.e.,

\[
x_t = d_t + \sum_{s=1}^{t} J_{s,t} f_s
\]

where \(d_t \triangleq x_0 + \sum_{i=1}^{t} c_i\) and \(J_{s,t} \triangleq \sum_{i=s}^{t} E_{s,i}\). As in Section 4.1.1, we implement the almost sure equality constraint \((21)\) via equality constraints on the policy parameters by setting \(d_T = 0\), and \(J_{t,T} = 0\) for all \(t\). We replace the almost sure inequality constraints \((22)–(23)\) with probabilistic relaxations, as in Section 4.1.2. With these assumptions, we compute the parameters of the optimal linear policy by solving the following stochastic program:

\[
\begin{aligned}
\max_{(E, c)} & \quad \mathbb{E} \left[ \sum_{t=1}^{T} \left( \left( d_t + \sum_{s=1}^{t} J_{s,t} f_s \right)^\top B f_t - \frac{1}{2} \left( c_t + \sum_{s=1}^{t} E_{s,t} f_s \right)^\top \Lambda \left( c_t + \sum_{s=1}^{t} E_{s,t} f_s \right) \right) \right] \\
\text{subject to} & \quad d_t = x_0 + \sum_{i=1}^{t} c_i, \quad 1 \leq t \leq T, \\
& \quad J_{s,t} = \sum_{i=s}^{t} E_{s,i}, \quad 1 \leq s \leq t \leq T, \\
& \quad \mathbb{P} \left( d_t + \sum_{s=1}^{t} J_{s,t} f_s < 0 \right) \leq \eta, \quad 1 \leq t \leq T, \\
& \quad \mathbb{P} \left( c_t + \sum_{s=1}^{t} E_{s,t} f_s > 0 \right) \leq \eta, \quad 1 \leq t \leq T, \\
& \quad d_T = 0, \\
& \quad J_{t,T} = 0, \quad 1 \leq t \leq T.
\end{aligned}
\]

Here, the parameter \(\eta \in (0, 1/2)\) controls the probability that the constraints \((22)–(23)\) are violated.\(^6\) Using the fact that the objective is an expectation of a quadratic expression in Gaussian random variables and the fact that the chance constraints can be handled using Lemma 1, \((29)\) can be explicitly written as a second-order cone program. This calculation is detailed in Appendix D. Then, \((29)\) can be solved using an off-the-shelf convex optimization solver.

The solution of the \((29)\) provides the desired linear policy, \(u_t = c_t + \sum_{s=1}^{t} E_{s,t} f_s\), in the return

\(^6\)We used the value \(\eta = 0.2\) in our simulation results.
predicting factors. However, due to the fact that some of the constraints of the original program in (24) are only probabilistically enforced, \( u_t \) may not be feasible for the original program. The projected optimal linear policy seeks a trade decision, \( \hat{u}_t \), which is the projection of \( u_t \) onto the constraint set (22)–(23), i.e.,

\[
\hat{u}_{i,t} = \max \left\{ -x_{i,t-1}, \min \{0, u_{i,t}\} \right\},
\]

for each time \( t < T \) and security \( i \).

5.3. Upper Bounds

In order to evaluate the quality of the policies described in Section 5.2, we compute a number of upper bounds on the performance of the any policy for the program (24), as follows:

- **Perfect foresight.** In this upper bound, we compute the value of an optimal policy with the perfect knowledge of future factor values. In particular, given a vector of factor realizations \( f \), consider the optimization problem

\[
\mathcal{V}_{PF}(f) \triangleq \max_u \sum_{t=1}^{T} \left( x_t^\top B f_t - \frac{1}{2} u_t^\top \Lambda u_t \right)
\]

subject to

\[
\begin{align*}
    u_t &= x_t - x_{t-1}, & t &= 1, \ldots, T, \\
    u_t &\leq 0, & x_t &\geq 0, & t &= 1, \ldots, T, \\
    x_T &= 0.
\end{align*}
\]

The value \( \mathcal{V}_{PF}(f) \) is the best that can be achieved with perfect foresight of a particular sample path of factors \( f \). Note that this can be readily computed by solving the quadratic program (30). Since the non-anticipating policies of the original program (24) are not able to utilize future factor information in making trading decisions, we have the upper bound

\[
\mathcal{V}^* \leq E[\mathcal{V}_{PF}(f)].
\]

This upper bound can be computed via Monte Carlo simulation over sample paths of factor realizations.

- **Unconstrained LQC.** The value of the LQC problem (27), where the inequality constraints (22)–(23) are relaxed, also provides an upper bound to (24). The expected value of the relaxed program can be exactly computed and yields the upper bound

\[
\mathcal{V}^* \leq \frac{1}{2} x_0^\top A_{xx,0} x_0 + \frac{1}{2} \left( \text{tr} \left( \Omega_0 (I - \Phi)^\top A_{ff,0} (I - \Phi) \right) + \sum_{t=0}^{T-2} \text{tr} (\Psi A_{ff,t}) \right),
\]

where the matrices \( A_{xx,0} \) and \( A_{ff,t} \) are derived in Appendix C.
Pathwise optimization. Given a sample path $f$ of factor realizations and a sequence $\zeta \triangleq (\zeta_1, \ldots, \zeta_T)$ of vectors $\zeta_t \in \mathbb{R}^K$ for each $t$, consider the quadratic optimization program

$$
\mathcal{V}^{PO}(f, \zeta) \triangleq \max_{u} \sum_{t=1}^{T} \left( x_t^\top B f_t - \zeta_t^\top \epsilon_t^{(1)} - \frac{1}{2} u_t^\top \Lambda u_t \right)
$$

subject to

$$
\begin{align*}
\epsilon_t^{(1)} &= f_t - (I - \Phi)f_{t-1}, & t &= 1, \ldots, T, \\
u_t &= x_t - x_{t-1}, & t &= 1, \ldots, T, \\
u_t &\leq 0, \, x_t \geq 0, & t &= 1, \ldots, T, \\
x_T &= 0.
\end{align*}
$$

(32)

It can be established (Desai et al., 2011; Brown and Smith, 2010) that for any $\zeta$, the upper bound

$$
\mathcal{V}^* \leq \mathbb{E}[\mathcal{V}^{PO}(f, \zeta)]
$$

holds — observe that the perfect foresight upper bound is a special case of this when $\zeta$ is zero. Roughly speaking, this upper bound corresponds to a relaxation of the non-anticipating policy requirement, and $\zeta$ correspond to a choice of Lagrange multipliers for this relaxation. The pathwise optimization upper bound corresponds to making a choice for $\zeta$ that results in an optimal upper bound, i.e.,

$$
\mathcal{V}^* \leq \min_{\zeta} \mathbb{E}[\mathcal{V}^{PO}(f, \zeta)].
$$

This minimization involves a convex objective function and can be computed via stochastic gradient descent; we refer the reader to Desai et al. (2011) for details.

5.4. Model Calibration

In this section, we describe the calibration the parameters of the optimal execution problem formulated in Section 5.1. We chose one of the most liquid stocks, Apple, Inc. (NASDAQ: AAPL), for our empirical study. We set the execution horizon to be 1 hour and trade intervals to be 5 minutes. Thus, setting a trade interval to be a one unit of time, we have a time horizon of $T = 12$.

We assume that the the initial position to be liquidated is $x_0 = 100,000$ shares.

In trade execution problems, the time horizon is typically a day, thus we will construct a factor model in the same time-frequency. We will use the intraday transaction prices of AAPL from the NYSE TAQ database on the trading days of January 4, 2010 (day 0) and January 5, 2010 (day 1) to construct $K = 2$ return predicting factors, each with a different mean reversion speed. We first divide each trading day into 78 time intervals, each 5 minutes in length. For each 5 minute interval, we calculate the average transaction price from all transactions in that interval. Let $p_t^{(d)}$ be the average price for interval $t = 1, \ldots, 78$ on day $d = 0, 1$. Let $f_{k,t}$ be the value of factor $k = 1, 2$ for interval $t = 2, \ldots, 78$, defined as follows

$$
f_{1,t} \triangleq p_t^{(1)} - p_{t-1}^{(1)}, \quad f_{2,t} \triangleq p_t^{(1)} - p_t^{(0)}.
$$
In other words, $f_{1,t}$ is the average price change over the previous 5 minute interval, while $f_{2,t}$ is the average price change relative to the previous day. Here, we can interpret the factors as the representations of value and momentum signals. Intuitively, the first factor can be considered as a ‘momentum’-type signal with fast mean reversion and the second factor as a ‘value’-type signal with slow mean reversion.

Given the price change of the security $r_{t+1} \triangleq p_{t+1}^{(1)} - p_{t}^{(1)}$, we can compute the estimate of the factor loading matrix, $B$, using the following pooled regression:

$$
\begin{align*}
 r_{t+1} &= 0.0726 + 0.3375 \ f_{1,t} - 0.0720 \ f_{2,t} + \epsilon_{t+1}, \\
(1.96) & \quad (3.11) & \quad (-2.2)
\end{align*}
$$

where the OLS t-statistics are reported in brackets. Thus,

$$
B = \begin{bmatrix}
0.3375 & -0.072
\end{bmatrix}.
$$

Similarly, we obtain the mean reversion rates for the factors,

$$
\begin{align*}
\Delta f_{1,t+1} &= -0.0353 \ f_{1,t} + \epsilon_{1,t+1}^{(1)}, \\
(-1.16) & \\
\Delta f_{2,t+1} &= -0.7146 \ f_{2,t} + \epsilon_{2,t+1}^{(1)}.
\end{align*}
$$

Thus,

$$
\Phi = \begin{bmatrix}
0.0353 & 0 \\
0 & 0.7146
\end{bmatrix}.
$$

The variance of the error terms is estimated to be

$$
\begin{align*}
\Sigma & \triangleq \text{Var}(\epsilon^{(1)}_t) = 0.0428, \\
\Psi & \triangleq \text{Var}(\epsilon^{(2)}_t) = \begin{bmatrix}
0.0378 & 0 \\
0 & 0.0947
\end{bmatrix}.
\end{align*}
$$

The distribution of the initial factor realization, $f_0$, is set to the stationary distribution under the given factor dynamics, i.e., $f_0$ is normally distributed with zero mean and covariance

$$
\Omega_0 \triangleq \sum_{t=1}^{\infty} (I - \Phi)^t \Psi (I - \Phi)^t = \begin{bmatrix}
0.0412 & 0 \\
0 & 1.3655
\end{bmatrix}.
$$

A rough estimate of the transaction cost coefficient $\Lambda = 2.14 \times 10^{-5}$ is used — this implies a transaction cost of $10 on a typical trade of 1,000 shares.

### 5.5. Numerical Results

Using the calibrated parameters from Section 5.4, we run a simulation with 50,000 trials to estimate the performance of each of the approximate policies of Section 5.2. In each trial, we sample the initial factor $f_0$, solve for the resulting policy of each approximate method, and compute its corresponding payoff. In order to evaluate the performance of each policy effectively, we use the same set of
## Table 1: Summary of the performance statistics of each policy, along with upper bounds. In the upper half of the table we consider the approximate policies. For each approximate policy, we divide the total payoff into two components, the alpha gains and the transaction costs. For each performance statistic, we report the mean value and the associated standard error. Finally, we report the average computation time (in seconds) for each policy per simulation trial. In the bottom half of the table, we report the computed upper bounds on the total payoff. For those methods which involve Monte Carlo simulation, standard errors are also reported.

<table>
<thead>
<tr>
<th>Policies</th>
<th>Alpha</th>
<th>TC</th>
<th>Total</th>
<th>CPU time (sec)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Deterministic</td>
<td>Mean 19.34</td>
<td>-15.81</td>
<td>3.53</td>
<td>0.82</td>
</tr>
<tr>
<td></td>
<td>S.E. 0.229</td>
<td>0.025</td>
<td>0.224</td>
<td></td>
</tr>
<tr>
<td>Model predictive control</td>
<td>Mean 21.25</td>
<td>-16.54</td>
<td>4.71</td>
<td>5.79</td>
</tr>
<tr>
<td></td>
<td>S.E. 0.233</td>
<td>0.023</td>
<td>0.225</td>
<td></td>
</tr>
<tr>
<td>Projected LQC</td>
<td>Mean 25.13</td>
<td>-19.40</td>
<td>5.73</td>
<td>0.02</td>
</tr>
<tr>
<td></td>
<td>S.E. 0.227</td>
<td>0.039</td>
<td>0.229</td>
<td></td>
</tr>
<tr>
<td>Optimal linear</td>
<td>Mean 23.24</td>
<td>-17.11</td>
<td>6.13</td>
<td>4.23</td>
</tr>
<tr>
<td></td>
<td>S.E. 0.233</td>
<td>0.025</td>
<td>0.224</td>
<td></td>
</tr>
<tr>
<td>Pathwise optimization</td>
<td>Mean 6.46</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>S.E. 0.04</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Perfect foresight</td>
<td>Mean 8.57</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>S.E. 0.223</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Unconstrained LQC</td>
<td>Mean 12.58</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>S.E. n/a</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

We observe that the optimal linear policy achieves the best performance. The gain of the optimal linear policy is approximately 7% over the next closest policy, which is the projected LQC policy. The performance of the other two policies is significantly worse. Since the projected LQC policy has a closed form expression (given a one time solution of recursive equations), its computation time per sample path is much smaller than that of the other policies, each of which involve solving at least one optimization problem per sample path. The remaining policies have roughly the same order of magnitude in computation time, with model predictive control (which solves a different optimization problem at every time step) having the longest running time.

Despite the higher total payoff for the optimal linear policy as compared to the projected LQC policy in Table 1, the relatively high standard errors preclude the immediate conclusion that...
the optimal linear policy achieves a statistically significant higher total payoff. Thus, in order to provide a more careful comparison, for each simulation trial, we consider the difference in alpha gains, transaction costs, and total payoff between these two policies. Table 2 shows the statistics of these differences, and establishes that the performance benefit of the optimal linear policy is statistically significant. Moreover, Table 2 reveals that the optimal linear policy achieves a better result by more carefully managing transaction costs, at the expense of not achieving the alpha gains of the projected LQC policy.

Finally, observe that the bottom half of Table 1 reports upper bounds on the total payoff of any policy, as computed using the methods described in Section 5.3. The pathwise optimization method achieves the tightest upper bound. Comparing this with the performance of the optimal linear policy, we conclude that the optimality gap of employing the optimal linear policy is less than 5% of the optimal value of the original program in (24).

6. Conclusion

This paper provides a highly tractable formulation for determining rebalancing rules in dynamic portfolio choice problems with involving complex models of return predictability. Our rebalancing rule is a linear function of past return predicting factors and can be utilized in a wide spectrum of portfolio choice models with realistic considerations for risk measures, transaction costs, and trading constraints. We illustrate the broad utility of our method by showing its applicability across a broad range of modeling assumptions on these portfolio optimization primitives. As long as the underlying dynamic portfolio optimization problem is a convex programming problem (i.e., concave objective and convex decision constraints), the modified optimization problem seeking the optimal parameters of the linear decision rule will be a convex programming problem that is tractable numerically. We demonstrate in an optimal execution problem that such modeling flexibility can offer significant practical benefits.

<table>
<thead>
<tr>
<th>(Optimal Linear) – (Projected LQC)</th>
<th>Alpha</th>
<th>TC</th>
<th>Total</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean</td>
<td>-1.89</td>
<td>2.29</td>
<td>0.40</td>
</tr>
<tr>
<td>S.E.</td>
<td>0.0137</td>
<td>0.0196</td>
<td>0.0095</td>
</tr>
</tbody>
</table>

Table 2: Detailed comparison between the alpha gains, transaction costs, and total performance of the optimal linear policy and projected dynamic policy. We observe that the standard error for the difference in total payoff is very small, thus, the performance gain by employing the optimal linear policy is statistically significant.
References


A. Proof of Lemma 1

Lemma 1. Given $\eta \in [0, 1/2]$, a non-zero vector $a \in \mathbb{R}^N$, and a scalar $b$, the chance constraint $P(a^T u_t > b) \leq \eta$ is equivalent to the second order cone constraint

$$a^T (c_t + M_t \theta_t) - b + \Phi^{-1}(1 - \eta) \left\| \Omega_t^{1/2} M_t^T a \right\|_2 \leq 0$$

on the policy coefficients $(c_t, M_t)$, where $\Phi^{-1}(\cdot)$ is the inverse cumulative normal distribution.

Proof. This proof follows standard arguments in convex optimization (see, e.g., Boyd and Vandenberghe, 2004). Let $\bar{u}_t$ and $V_t$ be the mean and the variance of $u_t$ as given in (13). Then,

$$P(a^T u_t > b) = P(\beta_t + \sigma_t Z > 0),$$

where

$$\beta_t \triangleq a^T \bar{u}_t - b, \quad \sigma_t \triangleq \|V_t^{1/2} a\|_2 \neq 0,$$

and $Z$ is a standard normal random variable. Thus,

$$P(a^T u_t > b) = 1 - \Phi(-\beta_t/\sigma_t).$$

Note that the this probability is less than or equal to $\eta$ if and only if

$$\beta_t + \Phi^{-1}(1 - \eta)\sigma_t \leq 0.$$

Substituting (13) into definitions for $\beta_t$ and $\sigma_t$, we obtain the desired result. ■

B. Exact Formulation of the Terminal Wealth Objective

Following the notation of Section 4, we will compute $E[W(x, r)^2]$ analytically and demonstrate that the resulting expression is a quadratic convex function of the policy coefficients. Without loss of generality, assume that $W_0 = 0$ and $\mu_t = 0$. Then, observe that

$$(33) \quad E \left[ W(x, r)^2 \right] = \sum_{t=1}^T \sum_{k=1}^T E \left[ (x_t^T r_{t+1})(x_k^T r_{k+1}) \right]$$

$$= \sum_{t=1}^T E \left[ x_t^T \epsilon_{t+1}^{(2)} \epsilon_{t+1}^{(2)\top} x_t \right] + \sum_{t=1}^T \sum_{k=1}^T E \left[ x_t^T B_f t f_k^\top B^\top x_k \right]$$

$$= \sum_{t=1}^T E \left[ x_t^T \epsilon_{t+1}^{(2)} \epsilon_{t+1}^{(2)\top} x_t \right] + \sum_{t=1}^T E \left[ x_t^T B_f t f_t^\top B^\top x_t \right] + 2 \sum_{t=1}^T \sum_{k=1}^T E \left[ x_t^T B_f t f_k^\top B^\top x_k \right].$$

We will consider each of the three terms in (33) separately.
The first term can be evaluated be first expectation, $E \left[ x_t^\top \epsilon_{t+1}^{(2)} (\epsilon_{t+1}^{(2)})^\top x_t \right]$, can be evaluated in the following form:

$$E \left[ x_t^\top \epsilon_{t+1}^{(2)} (\epsilon_{t+1}^{(2)})^\top x_t \right] = E \left[ x_t^\top E \left[ \epsilon_{t+1}^{(2)} (\epsilon_{t+1}^{(2)})^\top | F_t \right] x_t \right] = E \left[ x_t^\top \Sigma x_t \right] = (d_t + P_t\theta_t)^\top \Sigma (d_t + P_t\theta_t) + \text{tr} \left( \Sigma P_t\Omega_t P_t^\top \right).$$

Using the representation for $x_t$, we obtain

$$E \left[ x_t^\top B f_t f_t^\top B^\top x_t \right] = E \left[ \left( d_t + \sum_{s=1}^{t} J_{s,t} f_s \right)^\top B f_t f_t^\top B^\top \left( d_t + \sum_{s=1}^{t} J_{s,t} f_s \right) \right],$$

which can be simplified further with

$$E \left\{ \left( d_t + \sum_{s=1}^{t-1} J_{s,t} f_s + J_{t,t} (I - \Phi) f_{t-1} + \epsilon_t^{(1)} \right)^\top B (I - \Phi) f_{t-1} + \epsilon_t^{(1)} \right) \left( d_t + \sum_{s=1}^{t-1} J_{s,t} f_s + J_{t,t} (I - \Phi) f_{t-1} \right) \right. \left. B^\top \left( d_t + \sum_{s=1}^{t-1} J_{s,t} f_s + J_{t,t} (I - \Phi) f_{t-1} + \epsilon_t^{(1)} \right) \right\}.$$

The resulting computation deals with taking expectations of the product of jointly normal random variables with integer exponents:

$$E \left[ \left( d_t + \sum_{s=1}^{t-1} J_{s,t} f_s + J_{t,t} (I - \Phi) f_{t-1} \right)^\top B (I - \Phi) f_{t-1} (B (I - \Phi) f_{t-1})^\top \left( d_t + \sum_{s=1}^{t-1} J_{s,t} f_s + J_{t,t} (I - \Phi) f_{t-1} \right) \right] + E \left[ \left( \epsilon_t^{(1)} \right)^\top B (I - \Phi) f_{t-1} (B (I - \Phi) f_{t-1})^\top \epsilon_t^{(1)} \right] + E \left[ \left( \epsilon_t^{(1)} \right)^\top \epsilon_t^{(1)} \left( \epsilon_t^{(1)} \right)^\top \epsilon_t^{(1)} \right]$$

First define

$$\psi^2 \triangleq E \left[ \left( \epsilon_t^{(1)} \right)^\top \epsilon_t^{(1)} \left( \epsilon_t^{(1)} \right)^\top \epsilon_t^{(1)} \right] = \sum_{i=1}^{K} 3\Psi_{ii} + 2 \sum_{i=1}^{K} \sum_{j=1}^{i} \Psi_{ij},$$

where $\Psi_{ij}$ is the $(i,j)$th entry of the covariance matrix of the error terms for factor dynamics, $\Psi$. Using iterated expectations by conditioning on the information up to $t-1$, $F_{t-1}$, $E \left[ x_t^\top B f_t f_t^\top B^\top x_t \right]$. 

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\[ E \left[ \left( d_t + \sum_{s=1}^{t-1} J_s,t f_s + J_{t,t}(I - \Phi) f_{t-1} \right) B(I - \Phi) f_{t-1} (B(I - \Phi) f_{t-1})^\top \right] + \left( d_t + \bar{P}_{t-1} \theta_{t-1} \right) ^\top \Psi \left( d_t + \bar{P}_{t-1} \theta_{t-1} \right) + \text{tr}\left( \bar{P}_{t-1} \Omega_{t-1} \bar{P}_{t-1} ^\top \right) \text{tr}\left( (B(I - \Phi) \Omega_{t-1} ^{-1} (B(I - \Phi))^\top \Psi \right) } + \psi^2. \]

where

\begin{equation}
\bar{P}_{t-i} \triangleq \begin{bmatrix} J_{1,t} & J_{2,t} & \ldots & J_{t-i-1,t} & \bar{J}_{t-i,t} \end{bmatrix}, \quad \bar{J}_{t-i,t} \triangleq \left( \sum_{k=0}^{i} J_{t-k,t}(I - \Phi)^{i-k} \right). \end{equation}

We continue conditioning in a recursive fashion, and finally obtain the deterministic form

\[ E \left[ x_t ^\top B f_t f_t ^\top B ^\top x_t \right] = \left( d_t + \sum_{s=1}^{t} J_s,t(I - \Phi)^sf_0 \right) B(I - \Phi)^sf_0 \left( B(I - \Phi)^sf_0 \right) ^\top \left( d_t + \sum_{s=1}^{t} J_s,t(I - \Phi)^sf_0 \right) + \sum_{i=1}^{t} \left( d_t + \bar{P}_{t-i} \theta_{t-i} \right) ^\top \Psi \left( d_t + \bar{P}_{t-i} \theta_{t-i} \right) + \text{tr}\left( \bar{P}_{t-i} \Omega_{t-i} \bar{P}_{t-i} ^\top \right) \text{tr}\left( (B(I - \Phi)^i \Omega_{t-i} ^{-1} (B(I - \Phi)^i)^\top \Psi \right) + \psi^2, \]

which is convex quadratic function of the linear policy parameters.

The third expectation, \( E \left[ x_t ^\top B f_t f_k ^\top B ^\top x_k \right] \), can be computed using the same procedure.

### C. Derivation of the LQC Policies

We can derive a closed form solution for our trading policy when the problem satisfies the LQC framework. We guess a functional form for the value function and show that this functional form is preserved at each time step.

Using dynamic programming principle and \( u_t = (x_t - x_{t-1}) \), the value function \( V_t(x_{t-1}, f_t) \) satisfies

\[ V_{t-1}(x_{t-1}, f_t) = \text{maximize}_{x_t} \left( x_t ^\top (B f_t) - \frac{1}{2} (x_t - x_{t-1}) ^\top A(x_t - x_{t-1}) + E[V_t(x_t, f_{t+1})] \right). \]

We guess the following quadratic form for our value function:

\[ V_t(x_t, f_{t+1}) = -\frac{1}{2} x_t ^\top A_{xx,t} x_t + x_t ^\top A_{xf,t} f_{t+1} + \frac{1}{2} f_{t+1} ^\top A_{ff,t} f_{t+1} + \frac{1}{2} m_t. \]

Then,

\[ E[V_t(x_t, f_{t+1})] = -\frac{1}{2} x_t ^\top A_{xx,t} x_t + x_t ^\top A_{xf,t} (I - \Phi) f_t + \frac{1}{2} f_t ^\top (I - \Phi) ^\top A_{ff,t} (I - \Phi) f_t + \frac{1}{2} \text{tr}(\Psi A_{ff,t}) + m_t. \]
At the last period, we need \( x_T = 0 \), and our value function equals

\[
V_{T-1}(x_{T-1}, f_t) = -\frac{1}{2}x_{T-1}^\top \Lambda x_{T-1}
\]

which satisfies our functional form with

\[
A_{xx,T-1} = \Lambda \quad A_{xf,T-1} = 0 \quad A_{ff,T-1} = 0 \quad m_{T-1} = 0.
\]

For all \( t < T - 1 \), we maximize the quadratic objective

\[
-\frac{1}{2}x_t^\top Q_t x_t + x_t^\top q_t + b_t
\]

where

\[
Q_t = \Lambda + A_{xx,t},
\]

\[
q_t = \Lambda x_{t-1} + (B + A_{xf,t} (I - \Phi)) f_t,
\]

\[
b_t = \frac{1}{2} x_{t-1}^\top \Lambda x_{t-1} + \frac{1}{2} f_t^\top (I - \Phi)^\top A_{ff,t} (I - \Phi) f_t + \text{tr}(\Psi A_{ff,t}) + m_t.
\]

Then, the optimal \( x_t \) is given by

\[
Q_t^{-1} q_t + b_t
\]

and we obtain the following recursions:

\[
A_{xx,t-1} = -\Lambda (\Lambda + A_{xx,t})^{-1} \Lambda + \Lambda
\]

\[
A_{xf,t-1} = \Lambda (\Lambda + A_{xx,t})^{-1} (B + A_{xf,t} (I - \Phi))
\]

\[
A_{ff,t-1} = (B + A_{xf,t} (I - \Phi))^\top (\Lambda + A_{xx,t})^{-1} (B + A_{xf,t} (I - \Phi)) + (I - \Phi)^\top A_{ff,t} (I - \Phi)
\]

\[
m_{t-1} = \text{tr}(\Psi A_{ff,t}) + m_t.
\]

Using these recursions, we can compute the optimal expected payoff of the dynamic program. Using \( f_0 = N(0, \Omega_0) \),

\[
E[V_0(x_0, f_1)] = E[E[V_0(x_0, f_1)|f_0]]
\]

\[
= E \left[ -\frac{1}{2} x_0^\top A_{xx,0} x_0 + x_0^\top A_{xf,0} (I - \Phi) f_0 + \frac{1}{2} f_0^\top (I - \Phi)^\top A_{ff,0} (I - \Phi) f_0 + \frac{1}{2} (\text{tr}(\Omega_0 A_{ff,0}) + m_0) \right]
\]

\[
= -\frac{1}{2} x_0^\top A_{xx,0} x_0 + \frac{1}{2} \left( \text{tr} \left( \Omega_0 (I - \Phi)^\top A_{ff,0} (I - \Phi) \right) + \sum_{t=0}^{T-2} \text{tr}(\Psi A_{ff,t}) \right).
\]

**D. Exact Formulation of Best Linear Execution Policy**

We will first compute the expectation in the objective of (29) and write the equivalent deterministic form. We will then replace probabilistic constraints with deterministic constraints using Lemma 1, and finally obtain the deterministic version of the stochastic program in (29).
We start working with the expectation in the objective function. For each \( t \), we have to compute the expectation of the following two terms, \( E[x_t^T(Bf_t)] \) and \( E[u_t^T\Lambda u_t] \). First, we derive the statistics for \( f_t, u_t \) and \( x_t \). We first note that

\[
f_t = (I - \Phi)^t f_0 + \sum_{s=1}^{t} (I - \Phi)^{t-s} \epsilon_s^{(1)}.
\]

Letting \( F_t \triangleq (f_1, \ldots, f_t)^T \), then, in vector form, we have the following representation

\[
F_t = \begin{bmatrix}
(I - \Phi)f_0 \\
(I - \Phi)^2f_0 \\
\vdots \\
(I - \Phi)^{t-1}f_0 \\
(I - \Phi)^tf_0
\end{bmatrix} + \begin{bmatrix}
I & 0 & \cdots & 0 & 0 \\
(I - \Phi) & I & 0 & \cdots & 0 \\
\vdots & (I - \Phi) & \ddots & \cdots & 0 \\
(I - \Phi)^{t-1} & \cdots & (I - \Phi) & \ddots & 0 \\
(I - \Phi)^t & \cdots & (I - \Phi)^{t-1} & \cdots & (I - \Phi) \\
\end{bmatrix} \begin{bmatrix}
\epsilon_1^{(1)} \\
\epsilon_2^{(1)} \\
\vdots \\
\epsilon_{t-1}^{(1)} \\
\epsilon_t^{(1)}
\end{bmatrix}.
\]

Using this representation, we compute the mean

\[
\theta_t \triangleq \mathbb{E}[F_t] = \begin{bmatrix}
\delta_1 \\
\delta_2 \\
\vdots \\
\delta_{t-1} \\
\delta_t
\end{bmatrix} \triangleq \begin{bmatrix}
(I - \Phi)f_0 \\
(I - \Phi)^2f_0 \\
\vdots \\
(I - \Phi)^{t-1}f_0 \\
(I - \Phi)^tf_0
\end{bmatrix}
\]

and the covariance matrix

\[
\Omega_t \triangleq \text{Var}[F_t] = A_t \begin{bmatrix}
\Psi & 0 & \cdots & 0 & 0 \\
0 & \Psi & \cdots & 0 \\
\vdots & \ddots & \cdots & 0 \\
0 & \cdots & \Psi & 0 \\
0 & \cdots & 0 & \Psi
\end{bmatrix} A_t^T.
\]

Note that \( \Omega_t \) is a block diagonal matrix with \( t \) blocks of size \( K \times K \). Recall that in Section 4, we defined

\[
M_t \triangleq \begin{bmatrix}
E_{1,t} & E_{2,t} & \cdots & E_{t,t}
\end{bmatrix}
\]

Then, \( u_t = c_t + M_tF_t \) and we have the following moments for \( u_t \):

\[
\mu_t \triangleq \mathbb{E}[u_t] = c_t + M_t\theta_t
\]

\[
V_t \triangleq \text{Var}(u_t) = M_t\Omega_tM_t^T.
\]
Therefore, $u_t$ is normally distributed with mean $\mu_t$ and covariance matrix $V_t$. Similarly, we can obtain the statistics for $x_t$. Using (15),

$$
\kappa_t \triangleq \mathbb{E}[x_t] = d_t + P_t \theta_t \\
Y_t \triangleq \text{Var}(x_t) = P_t \Omega_t P_t^T.
$$

We note the following fact from multivariate statistics.

**Fact 1.** If $z$ is a random vector with mean $\mu$ and variance $\Sigma$, and $Q$ is positive definite matrix, then

$$
\mathbb{E}[z^T Q z] = \text{tr}(Q \Sigma) + \mu^T Q \mu.
$$

Using Fact 1 we can compute each term in the expectation.

$$
\mathbb{E} \left[ x_t^T (B f_t) \right] = \mathbb{E} \left[ d_t^T B f_t + \sum_{s=1}^t f_s^T J_{s,t}^T B f_t \right] \\
= d_t^T B \delta_t + \sum_{s=1}^t \mathbb{E} \left[ f_s^T J_{s,t}^T B \mathbb{E}[f_t|f_s] \right] \\
= d_t^T B \delta_t + \sum_{s=1}^t \mathbb{E} \left[ f_s^T J_{s,t}^T B (I - \Phi)^{t-s} f_s \right] \\
= d_t^T B \delta_t + \sum_{s=1}^t \left( \delta_s^T (B (I - \Phi)^{t-s} J_{s,t}^T) \delta_s + \text{tr} \left( B (I - \Phi)^{t-s} J_{s,t}^T \omega_s \right) \right)
$$

where $\omega_s$ is the $s$th diagonal block matrix of $\Omega_t$ having a size of $K \times K$. Finally, for the transaction cost terms,

$$
\mathbb{E} \left[ u_t^T \Lambda u_t \right] = \mathbb{E} \left[ (c_t + M_t \theta_t)^T \Lambda (c_t + M_t \theta_t) \right] \\
= (c_t + M_t \theta_t)^T \Lambda (c_t + M_t \theta_t) + \text{tr} \left( \Lambda M_t \Omega_t M_t^T \right)
$$

Summing up all the terms, our final objective function in deterministic form equals

$$
\maximize_{c_t, E_{s,t}} \sum_{t=1}^T \left\{ d_t^T B \delta_t + \sum_{s=1}^t \left( \delta_s^T (B (I - \Phi)^{t-s} J_{s,t}^T) \delta_s + \text{tr} \left( B (I - \Phi)^{t-s} E_{s,t}^T \omega_s \right) \right) \\
+ \frac{1}{2} \left( (c_t + M_t \theta_t)^T \Lambda (c_t + M_t \theta_t) + \text{tr} \left( \Lambda M_t \Omega_t M_t^T \right) \right) \right\},
$$

which is a quadratic function of the policy parameters.

We now rewrite the equality constraint, $x_T = 0$ in terms of policy parameters. In order to enforce this equality, we need

$$d_T = 0 \text{ and } J_{s,T} = 0 \quad s = 1, \ldots, T.$$
Lastly, we replace probabilistic constraints with deterministic constraints using Lemma 1. Note that 
\( \Pr(x_t \leq 0) \leq \eta \) can be written as \( \Pr(-x_t \geq 0) \leq \eta \). Then, using Lemma 1,
\[
(-d_t - P_t \theta_t) + \Phi^{-1}(1 - \eta) \left\| \left( P_t \Omega_t P_t^\top \right)^{1/2} \right\|_2 \leq 0.
\]

Similarly, we obtain that \( \Pr(u_t \geq 0) \leq \eta \) can be replaced by
\[
(c_t + M_t \theta_t) + \Phi^{-1}(1 - \eta) \left\| \left( M_t \Omega_t M_t^\top \right)^{1/2} \right\|_2 \leq 0.
\]

Combining all the results, we obtain the deterministic version of the stochastic program in (29), a second-order cone program:

\[
\begin{align*}
\text{maximize} & \quad \sum_{t=1}^{T} \left \{ d_t^\top B \delta_t + \sum_{s=1}^{t} \left( \delta_s^\top (B(I - \Phi)^{t-s} J_{s,t}^\top) \delta_s + \text{tr} \left( B(I - \Phi)^{t-s} E_{s,t}^\top \omega_s \right) \right) \right. \\
& \quad \left. + \frac{1}{2} \left( (c_t + M_t \theta_t)^\top \Lambda (c_t + M_t \theta_t) + \text{tr} \left( M_t^\top \Lambda M_t \Omega_t \right) \right) \right \} \\
\text{subject to} & \quad d_t = x_0 + \sum_{i=1}^{t} c_i \quad t = 1, \ldots, T, \\
& \quad J_{s,t} = \sum_{i=s}^{t} E_{s,i} \quad 1 \leq s \leq t \leq T, \\
& \quad (-d_t - P_t \theta_t) + \Phi^{-1}(1 - \eta) \left\| \left( P_t \Omega_t P_t^\top \right)^{1/2} \right\|_2 \leq 0 \quad t = 1, \ldots, T, \\
& \quad (c_t + M_t \theta_t) + \Phi^{-1}(1 - \eta) \left\| \left( M_t \Omega_t M_t^\top \right)^{1/2} \right\|_2 \leq 0 \quad t = 1, \ldots, T, \\
& \quad d_T = 0 \text{ and } J_{s,T} = 0.
\end{align*}
\]

Note that the number of decision variables is considerably greater than that of the original execution problem in (24). Total number of decision variables in a problem with \( N \) securities, \( K \) factors and \( T \) periods equals \( 2NT + NK(T + 1) \) which is on the order of \( O(NKT^2) \).